On the size of the block of 1 for Ξ -coalescents with dust

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Abstract We study the frequency process f_1 of the block of 1 for a Ξ -coalescent Π with dust. If Π stays infinite, f_1 is a jump-hold process which can be expressed as a sum of broken parts from a stick-breaking procedure with uncorrelated, but in general non-independent, stick lengths with common mean. For Dirac- Λ -coalescents with $\Lambda = \delta_p$, $p \in [\frac{1}{2}, 1)$, f_1 is not Markovian, whereas its jump chain is Markovian. For simple Λ -coalescents the distribution of f_1 at its first jump, the asymptotic frequency of the minimal clade of 1, is expressed via conditionally independent shifted geometric distributions.

Keywords *Ξ*-coalescent, coalescent with dust, Poisson point process, minimal clade, exchangeability

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1 Introduction and results

Independently introduced in [33] and [30], Ξ -coalescents are exchangeable Markovian processes $\Pi = (\Pi_t)_{t\geq 0}$ on the set of partitions of $\mathbb{N} := \{1, 2, ...\}$ whose transitions are due to mergers of partition blocks. The distribution of Π is characterised by

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a finite measure Ξ on the infinite simplex

$$\Delta := \{ x = (x_1, x_2, \ldots) : x_1 \ge x_2 \ge \cdots \ge 0, |x| \le 1 \},\$$

where $|x| := \sum_{i \in \mathbb{N}} x_i$. We exclude $\Xi = 0$, since it leads to a coalescent without coalescence events. Ξ -coalescents allow that disjoint subsets of blocks merge into distinct new blocks, hence they are also called coalescents with simultaneous multiple mergers. If Ξ is concentrated on $[0, 1] \times \{0\} \times \{0\} \times \cdots$, only a single set of blocks is allowed to merge. Such a coalescent is a Λ -coalescent, see [32]. In this case, Λ is a finite measure on [0, 1], the restriction of Ξ on the first coordinate of Δ . The restriction $\Pi^{(n)}$ of Π on $[n] := \{1, \ldots, n\}$ is called the Ξ -n-coalescent. Denote the blocks of Π_t by $(B_i(t))_{i \in \mathbb{N}}$, where *i* is the least element of the block (we set $B_i(t) = \emptyset$ if *i* is not a least element of a block). Clearly, $1 \in B_1(t)$. We call $B_1(t)$ the block of 1 at time *t*. Due to the exchangeability of the Ξ -coalescent, Kingman's correspondence ensures that, for every $t \ge 0$, the asymptotic frequencies

$$f_i(t) := \lim_{n \to \infty} \frac{|B_i(t) \cap [n]|}{n}, \quad i \in \mathbb{N},$$
(1)

exist almost surely, where |A| denotes the cardinality of the set A.

The family of Ξ -coalescents is a diverse class of processes with very different properties, see e.g. the review [15] for Λ -coalescents. We will focus on Ξ -coalescents with dust, i.e. Ξ fulfils (see [33])

$$\mu_{-1} := \int_{\Delta} |x| \nu_0(dx) < \infty, \tag{2}$$

where $v_0(dx) = \Xi(dx)/(x, x)$ with $(x, x) := \sum_{i \in \mathbb{N}} x_i^2$ for $x = (x_1, x_2, ...) \in \Delta$. These coalescents are characterised by a non-zero probability that, at any time *t*, there is a positive fraction of \mathbb{N} , the dust, that has not yet merged. Note that $i \in \mathbb{N}$ is part of the dust at time *t* if and only if $\{i\}$ is a block at time *t*, which is called a singleton block. The asymptotic frequency of the dust component is $S_t := 1 - \sum_{i \in \mathbb{N}} f_i(t)$. Having dust is equivalent to $P(S_t > 0) > 0$ for all t > 0. We are interested in Ξ coalescents which stay infinite, i.e. which almost surely have an infinite number of blocks for each t > 0. We will put some further emphasis on simple Λ -coalescents satisfying

$$\mu_{-2} := \int_{[0,1]} x^{-2} \Lambda(dx) < \infty.$$
(3)

This class includes Dirac coalescents with $\Lambda = \delta_p$, the Dirac measure in $p \in (0, 1]$. Consider the frequency process $f_1 := (f_1(t))_{t \ge 0}$ of the block of 1. For Λ -coalescents, Pitman characterises f_1 as follows (reproduced from [32], adjusted to our notation).

Proposition 1. [32, Proposition 30] No matter what Λ , the process f_1 is an increasing pure jump process with càdlàg paths, $f_1(0) = 0$ and $\lim_{t\to\infty} f_1(t) = 1$. If $\mu_{-1} = \infty$ then almost surely $f_1(t) > 0$ for all t > 0 and $\lim_{t\to0} f_1(t) = 0$. If $\mu_{-1} < \infty$ then f_1 starts by holding at zero until an exponential time with rate μ_{-1} , when it enters (0, 1] by a jump, and proceeds thereafter by a succession of holds and jumps, with holding rates bounded above by μ_{-1} . Moreover, in [32, Section 3.9], a general formula for the moments of $f_1(t)$ for fixed t > 0 is provided.

For two particular coalescents without dust, further properties of f_1 are known. For Kingman's *n*-coalescent ($\Lambda = \delta_0$), the complete distribution of block sizes is explicitly known, see [24, Theorem 1], from which one can derive some properties of the block of 1 due to exchangeability. For the Bolthausen–Sznitman coalescent (Λ the uniform distribution on [0, 1]) the block of 1 can be characterised as in [32, Corollary 16]. For instance, f_1 is Markovian for the Bolthausen–Sznitman coalescent.

Different specific aspects of the block of 1 have been analysed for different Λ/Ξ *n*-coalescents including their asymptotics for $n \to \infty$.

- External branch length: The waiting time for the first jump of the block of 1 in the *n*-coalescent, see e.g. [6–8, 13, 22, 28].
- Minimal clade size: The size M_n of the block of 1 for the *n*-coalescent at its first jump. For Kingman's *n*-coalescent and for Λ beta-distributed with parameters (2-α, α) with α ∈ (1, 2), X_n converges in distribution for n → ∞, see [6] and [34]. For the Bolthausen–Sznitman *n*-coalescent, log(M_n)/log(n) converges in distribution [14]. These results do not cover Λ/Ξ-coalescents with dust.
- The number of blocks involved in the first merger of the block of 1, see [34]. The results cover *A*-coalescents with dust.
- The number of blocks involved in the last merger of the block of 1, see [1, 2, 19, 17, 23, 29].
- The small-time behaviour of the block of 1, see [5, 34].

Due to the exchangeability of the Ξ -coalescent, any result for the distribution of the block of 1 holds true for the block containing any other $i \in \mathbb{N}$. We want to further describe f_1 for Ξ -coalescents with dust. For any finite measure Ξ on Δ which fulfils (2), we introduce

$$\gamma := \frac{\Xi(\Delta)}{\mu_{-1}}.$$
(4)

We see that $\gamma \in (0, 1]$, since

$$0 < \Xi(\Delta) = \int_{\Delta} (x, x) \nu_0(dx) \le \int_{\Delta} |x| \nu_0(dx) = \mu_{-1} < \infty.$$

Define $\Delta_f := \bigcup_{k \in \mathbb{N}} \{x \in \Delta : x_1 + \dots + x_k = 1\}$. We extend Proposition 1 for Ξ -coalescents with dust which stay infinite, i.e. have almost surely infinitely many blocks for each $t \ge 0$ (equivalent to $\Xi(\Delta_f) = 0$, see Lemma 4). While the extension to Ξ -coalescents and the explicit waiting time distributions are a direct follow-up from Pitman's proof, we provide a more detailed description of the jump heights of f_1 . Proposition 1 ensures that the jumps of f_1 are separated by (almost surely) positive waiting times, we denote the value of f_1 at its *k*th jump with $f_1[k]$ for $k \in \mathbb{N}$.

Theorem 1. In any Ξ -coalescent Π with dust and $\Xi(\Delta_f) = 0$, the asymptotic frequency process $f_1 := (f_1(t))_{t \ge 0}$ of the block of 1, defined by Eq. (1), is an increasing pure jump process with càdlàg paths, $f_1(0) = 0$ and $\lim_{t\to\infty} f_1(t) = 1$,

but $f_1(t) < 1$ for t > 0 almost surely. The waiting times between almost surely infinitely many jumps are distributed as independent $\text{Exp}(\mu_{-1})$ random variables. Its jump chain $(f_1[k])_{k\in\mathbb{N}}$ can be expressed via stick-breaking

$$f_1[k] = \sum_{i=1}^k X_i \prod_{j=1}^{i-1} (1 - X_j),$$
(5)

where $(X_j)_{j \in \mathbb{N}}$ are pairwise uncorrelated, $X_j > 0$ almost surely and $E(X_j) = \gamma$ for all $j \in \mathbb{N}$. In particular, $E(f_1[k]) = 1 - (1 - \gamma)^k$. In general, $(X_j)_{j \in \mathbb{N}}$ are neither independent nor identically distributed.

Remark 1. From Theorem 1, the dependence between f_1 and its jump times is readily seen as follows. Recall [32, Eq. (51)] that $E(f_1(t)) = 1 - e^{-t}$ for any Λ -coalescent with $\Lambda([0, 1]) = 1$. If we would have independence, integrating $E(f_1(t))$ over the waiting time distribution $Exp(\mu_{-1})$ for the first jump of f_1 would yield $E(f_1[1]) =$ $(1 + \mu_{-1})^{-1}$, in contradiction to $E(f_1[1]) = 1/\mu_{-1}$ by Theorem 1.

Dirac coalescents ($\Lambda = \delta_p$ for some $p \in (0, 1]$) are a family of Λ -coalescents with dust. They have been introduced as simplified models for populations in species with skewed offspring distributions (reproduction sweepstakes), see [9]. Their jump chains (discrete time Dirac coalescents) can also arise as large population size limits in conditional branching process models [21, Theorem 2.5].

We further characterise f_1 as follows, including an explicit formula for its distribution at its first jump.

Proposition 2. Let $\Lambda = \delta_p$, $p \in [\frac{1}{2}, 1)$ and q := 1 - p. f_1 takes values in the set

$$\mathcal{M}_{p} := \left\{ \sum_{i \in \mathbb{N}} b_{i} p q^{i-1} : b_{i} \in \{0, 1\}, 1 \le \sum_{i \in \mathbb{N}} b_{i} < \infty \right\}.$$
(6)

For $x = \sum_{i \in \mathbb{N}} b_i p q^{i-1} \in \mathcal{M}_p$, we have

$$P(f_1[1] = x) = pq^{j-1} \prod_{i \in J \setminus \{j\}} P(Y + i \in J) \prod_{i \in [j-1] \setminus J} P(Y + i \notin J) > 0, \quad (7)$$

where $Y \stackrel{d}{=} Geo(p)$, $J := \{i \in \mathbb{N} | b_i = 1\}$ and $j := \max J$. The process f_1 is not Markovian whereas its jump chain $(f_1[k])_{k \in \mathbb{N}}$ is Markovian.

Remarks 2.

- The law of $f_1[1]$ is a discrete measure on [0, 1] for Dirac coalescents. Surprisingly different properties arise for different values of p. For instance, $\mathcal{M}_{2/3} = \{\sum_{i \in \mathbb{N}} b_i 3^{-i} : b_i \in \{0, 2\}, 1 \leq \sum_{i \in \mathbb{N}} b_i < \infty\}$ is a subset of the ternary Cantor set which is nowhere dense in [0, 1], whereas $\mathcal{M}_{1/2}$, the set of all $x \in [0, 1]$ with finite 2-adic expansion, is dense in [0, 1].
- We omitted $f_1[1]$ for the star-shaped coalescent ($\Lambda = \delta_1$), since it just jumps from 0 to 1 at time $T \stackrel{d}{=} \text{Exp}(1)$.

 Recall that f₁ is Markovian for the Bolthausen–Sznitman coalescent in contrast to f₁ for the Dirac coalescents specified above.

Our key motivation was to provide a more detailed description of the jump chain of f_1 , especially properties of the value $f_1[1]$ at the first jump which is the asymptotic frequency of the minimal clade. Theorem 1 provides a first-order limit result for all Ξ -coalescents with dust.

Corollary 1. Let Π be a Ξ -coalescent with dust and $\Pi^{(n)}$ its restriction on [n]. Let M_n be the minimal clade size, i.e. the size of the block of 1 at its first merger in $\Pi^{(n)}$. Then, $M_n/n \to f_1[1]$ almost surely, $f_1[1] > 0$ almost surely and $E(f_1[1]) = \gamma$.

Compared to the known results listed above for the minimal clade size for dustfree coalescents, the minimal clade size is much larger asymptotically for $n \to \infty$ (O(n) compared to o(n)).

The law of $f_1[1]$ in (7) follows from the following more general description of $f_1[1]$ for simple Λ -coalescents. We introduce, for a finite measure Λ on [0, 1] with $\mu_{-1} = \int_0^1 x^{-1} \Lambda(dx) < \infty$,

$$\alpha := \frac{\mu_{-1}}{\mu_{-2}} = \frac{\int_0^1 x^{-1} \Lambda(dx)}{\int_0^1 x^{-2} \Lambda(dx)}.$$
(8)

We have $\alpha \in [0, 1]$ since $x^{-1} \le x^{-2}$ on (0, 1] (if $\mu_{-1} < \infty$, we have $\Lambda(\{0\}) = 0$). Additionally, $\alpha > 0$ if and only if $\mu_{-2} < \infty$, so if Λ characterises a simple coalescent (recall that $\mu_{-2} \ge \mu_{-1} > 0$ since we exclude $\Lambda = 0$).

Proposition 3. Let A fulfil (3). Then,

$$f_1[1] = \sum_{i=1}^{C} B_i^{(C)} P_i \prod_{j \in [i-1]} (1 - P_j) = \sum_{i \in \mathbb{N}} P_i \prod_{j \in [i-1]} (1 - P_j) \sum_{k \ge i} B_i^{(k)} \mathbb{1}_{\{C=k\}}, \quad (9)$$

where $(P_i)_{i \in \mathbb{N}}$ are *i.i.d.* with $P_i \stackrel{d}{=} \mu_{-2}^{-1} x^{-2} \Lambda(dx)$. We have

$$P(C = k | (P_i)_{i \in \mathbb{N}}) = P_k \prod_{j \in [k-1]} (1 - P_j), C \text{ is } Geo(\alpha) \text{-distributed.}$$

Given $(P_i)_{i \in \mathbb{N}}$, C and $(B_i^{(k)})_{k \in \mathbb{N}, i \in [k]}$ are independent and $(B_i^{(k)})_{k \in \mathbb{N}, i \in [k]}$ is defined as

$$P\left(\left(B_{1}^{(j)},\ldots,B_{j}^{(j)}\right)=b|(P_{i})_{i\in\mathbb{N}}\right)$$
$$=\prod_{i\in J\setminus\{j\}}P\left(I(i)\in J|(P_{i})_{i\in\mathbb{N}}\right)\prod_{i\in[j-1]\setminus J}P\left(I(i)\notin J|(P_{i})_{i\in\mathbb{N}}\right) almost surely,$$
(10)

where $b = (b_1, \ldots, b_j) \in \{0, 1\}^{j-1} \times \{1\}$, $J := \{i \in [j] | b_i = 1\}$ and, for each $i \in \mathbb{N}$, $I(i) := \min\{j \ge i+1 : B_i^{(j)} = 1\}$. We have

(i)
$$P(I(i) = i + k | (P_i)_{i \in \mathbb{N}}) = P_{i+k} \prod_{l=i+1}^{i+k-1} (1 - P_l)$$
 almost surely for $k \in \mathbb{N}$.

(ii) For any $i \in \mathbb{N}$, I(i) - i is $Geo(\alpha)$ -distributed on \mathbb{N} . Given $(P_i)_{i \in \mathbb{N}}$, $(I(i))_{i \in \mathbb{N}}$ are independent.

Remarks 3.

- The distribution of C is known from [16, Proposition 3.1].
- The distribution of $f_1[1]$ for Dirac coalescents with $p > \frac{1}{2}$ has a structure somewhat similar to the Cantor distribution, see e.g. [26] and [18]. The Cantor distribution is the law of $\sum_{i \in \mathbb{N}} B_i pq^{i-1}$ for $p \in (0, 1)$, where $(B_i)_{i \in \mathbb{N}}$ are i.i.d. Bernoulli variables with success probability $\frac{1}{2}$, whereas in our case $(B_i)_{i \in \mathbb{N}}$ are dependent Bernoulli variables with success probability p_i^{1} , whereas in our case $(B_i)_{i \in \mathbb{N}}$ are dependent Bernoulli variables with success probabilities $P(B_i = 1) = P(\sum_{k \ge i} B_i^{(k)} \mathbb{1}_{\{C=k\}} = 1) = pq^{i-1} + \sum_{k>i} p^2 q^{k-1} = pq^{i-1}(1+q)$, see Eq. (9). The Cantor distribution is a shifted infinite Bernoulli convolution. Infinite Bernoulli convolutions are the set of distributions of $\sum_{i \in \mathbb{N}} \omega_i(-1)^{B_i}$ with $\omega_i \in \mathbb{R}$ for $i \in \mathbb{N}$ satisfying $\sum_{i \in \mathbb{N}} \omega_i^2 < \infty$, see [31, Section 2]. They have been an active field of research since the 1930's, e.g. see [10, 35] and the survey [31].

Our main tool for the proofs is Schweinsberg's Poisson construction of the Ξ coalescent. The article is organised as follows. We recall (properties of) the Poisson
construction in Section 2. Section 3 characterises staying infinite for Ξ -coalescents
with dust. These prerequisites are then used to prove the results for Ξ -coalescents
with dust in Section 4 and for simple Λ -coalescents in Section 5.

2 Poisson construction of a Ξ -coalescent and the block of 1

We recall the construction of a Ξ -*n*-coalescent Π from [33]. We are only interested in constructing a Ξ -coalescent with dust, which implies $\Xi(\{0\}) = 0$, see Eq. (2).

Let \mathcal{P} be a Poisson point process on $A = [0, \infty) \times \mathbb{N}_0^\infty$ with intensity measure

$$\nu = dt \otimes \int_{\Delta} \otimes_{n \in \mathbb{N}} P^{(x)} \nu_0(dx), \tag{11}$$

where, for $x \in \Delta$, $P^{(x)}$ is a probability measure on \mathbb{N}_0 with $P^{(x)}(\{k\}) = x_k$ and $P^{(x)}(\{0\}) = 1 - |x|$ (Kingman's paintbox) and v_0 is defined as in Eq. (2). For $n \in \mathbb{N}$, the restriction $\Pi^{(n)}$ of Π to [n] can be constructed by starting at t = 0 with each $i \in [n]$ in its own block. Then, for each subsequent time (T =)t with a Poisson point $(T, (K_i)_{i \in \mathbb{N}})$, merge all present blocks i (at most n) with identical $k_i > 0$, where i is the least element of the block (there are only finitely many points of \mathcal{P} that lead to a merger of blocks in [n]). Π is then pathwise defined by its restrictions $(\Pi^{(n)})_{n \in \mathbb{N}}$. From now on we will assume without loss of generality that the Ξ -coalescent with dust is constructed via the Poisson process \mathcal{P} .

The block of 1 can only merge at Poisson points $P = (T, (K_i)_{i \in \mathbb{N}})$ with $K_1 > 0$. We take a closer look at these Poisson points. We introduce exchangeable(Q) indicators following [32, p.1884]: These are exchangeable Bernoulli variables which are conditionally i.i.d. given a random variable X with distribution Q on [0, 1] which gives their success probability. Alternatively, we denote these as exchangeable(X) indicators if we can specify X.

Lemma 1. For any finite measure Ξ on Δ fulfilling (2), \mathcal{P} splits into two independent Poisson processes

$$\mathcal{P}_1 := \{ (T, (K_i)_{i \in \mathbb{N}}) \in \mathcal{P} : K_1 > 0 \} \text{ and } \mathcal{P}_2 := \{ (T, (K_i)_{i \in \mathbb{N}}) \in \mathcal{P} : K_1 = 0 \}.$$

 \mathcal{P}_1 has almost surely finitely many points on any set $[0, t] \times \mathbb{N}_0^{\infty}$, thus we can order

$$\mathcal{P}_1 = \left(\left(T_j, \left(K_i^{(j)} \right)_{i \in \mathbb{N}} \right) \right)_{j \in \mathbb{N}^3}$$

where $T_j < T_{j+1}$ almost surely for $j \in \mathbb{N}$.

 $(T_j)_{j \in \mathbb{N}}$ is a homogeneous Poisson process on $[0, \infty)$ with intensity μ_{-1} .

 $((K_i^{(j)})_{i \in \mathbb{N}})_{j \in \mathbb{N}}$ is an i.i.d. sequence in j and $(1_{\{K_1^{(1)}=K_i^{(1)}\}})_{i\geq 2}$ are exchangeable(Q) indicators with

$$Q := \frac{1}{\mu_{-1}} \int_{\Delta} \sum_{i \in \mathbb{N}} x_i \delta_{x_i} v_0(dx),$$

which is a probability measure on [0, 1]. For $X \stackrel{d}{=} Q$, we have X > 0 almost surely and $E(X) = \gamma$.

Proof. \mathcal{P}_1 and \mathcal{P}_2 are obtained by restricting \mathcal{P} on the disjoint subsets $A_1 := [0, \infty) \times \mathbb{N} \times \mathbb{N}_0^\infty$ and $A_2 := [0, \infty) \times \{0\} \times \mathbb{N}_0^\infty$ of A. Thus, \mathcal{P}_1 and \mathcal{P}_2 are independent Poisson processes (restriction theorem [25, p.17]) with intensity measures $v_1 = v(\cdot \cap A_1)$ and $v_2 = v(\cdot \cap A_2)$. For any Borel set $B \subseteq [0, \infty)$ and λ being the Lebesgue measure,

$$\nu_1\left(B \times \mathbb{N}_0^\infty\right) = \lambda(B) \int_\Delta \underbrace{P^{(x)}(\mathbb{N})}_{=|x|} \prod_{i \ge 2} \underbrace{P^{(x)}(\mathbb{N}_0)}_{=1} \nu_0(dx) = \lambda(B)\mu_{-1}.$$
(12)

Thus, on any bounded set B, \mathcal{P}_1 has almost surely finitely many points, which can be ordered as described. Projecting \mathcal{P}_1 on the first coordinate *t* of *A* yields a Poisson process with intensity measure $\mu_{-1}dt$ (mapping theorem [25, p.18]).

Now, we project the points of \mathcal{P}_1 on the coordinate of $(K_i^{(j)})_{i \in \mathbb{N}}$. Recall the construction of a Poisson process as a collection of i.i.d. variables with distribution $(\mu(C))^{-1}\mu$ on sets of finite mass *C* of the intensity measure μ , e.g. [25, p.23]. It shows that we can treat the collection of $(T_j, (K_i^{(j)})_{i \in \mathbb{N}})$ with, for instance, $T_i \in [k, k + 1)$ for $k \in \mathbb{N}$ as a random number of i.i.d. random variables with distribution $(1 \cdot \mu_{-1})^{-1}\nu_1$. Since ν_1 has a product structure on A_1 , we have that $((K_i^{(j)})_{i \in \mathbb{N}})_{j \in \mathbb{N}}$ are i.i.d. in *j* and have distribution, for $m \in \mathbb{N}$,

$$P((K_i^{(1)} = l_i)_{i \in [m]}) = \frac{1}{\mu_{-1}} \int_{\Delta} \prod_{i \in [m]} P^{(x)}(l_i) \nu_0(dx) = \frac{1}{\mu_{-1}} \int_{\Delta} \prod_{i \in [m]} x_{l_i} \nu_0(dx)$$
(13)

for $l_1 \in \mathbb{N}$ and $l_2, \ldots, l_m \in \mathbb{N}_0$. Consider $(1_{\{K_1^{(1)}=K_i^{(1)}\}})_{i\geq 2}$. To show that they are exchangeable(Q) indicators, [32, Eq. (27)] has to be fulfilled, i.e. we need to show $P(\{i \in [m] : K_i^{(1)} = K_1^{(1)}\} = M) = E(X^{|M|-1}(1-X)^{m-|M|})$ for $X \stackrel{d}{=} Q$ and any $M \subseteq [m]$ with $1 \in M$. Using Eq. (13) we compute

$$P(\left\{i \in [m]: K_i^{(1)} = K_1^{(1)}\right\} = M) = \sum_{j \in \mathbb{N}} P(\left\{i \in [m]: K_i^{(1)} = j\right\} = M, K_1^{(1)} = j)$$

$$= \frac{1}{\mu_{-1}} \int_{\Delta} \sum_{j \in \mathbb{N}} x_j^{|M|} (1 - x_j)^{m - |M|} v_0(dx)$$
$$= E \left(X^{|M| - 1} (1 - X)^{m - |M|} \right).$$

Clearly, P(X > 0) = 1 since $\Xi(\{0\}) = 0$ and $E(X) = \mu_{-1}^{-1} \int_{\Delta} (x, x) v_0(dx) = \gamma$.

Remarks 4.

- The properties of the exchangeable(Q) indicators remind of [32, Lemma 21, Theorem 4] and [33, Proposition 6]. Restricting \mathcal{P} to points with $K_1 = K_2 > 0$ we can reproduce their results analogously to the proof of Lemma 1.
- Q can be seen as the expected value of the random probability measure $Q_x := |x|^{-1} \sum_{i \in \mathbb{N}} x_i \delta_{x_i}$ for $x \in \Delta$ with x drawn from $\mu_{-1}^{-1} |x| v_0(dx)$. In the Poisson construction, this means we draw a "paintbox" $x \in \Delta$ and then record in which box the ball of 1 falls, if we only allow it to fall in boxes 1, 2, ...
- Consider a simple Λ -coalescent. Projecting \mathcal{P}_2 on its first component, so $(T, (K_i)_{i \in \mathbb{N}}) \mapsto T$, yields a homogeneous Poisson process with intensity $\mu_{-2} \mu_{-1} < \infty$. To see this, proceed analogously as for \mathcal{P}_1 . Then, Eq. (12) for v_2 reads the same except for replacing $P^{(x)}(\mathbb{N})$ by $P^{(x)}(\{0\}) = 1 |x|$.

For a Λ -coalescent (with $\Lambda(\{0\}) = 0$) the Poisson construction simplifies, since Ξ only has mass on $\{x \in \Delta : x_2 = x_3 = \cdots = 0\}$ and thus \mathcal{P} can be seen as a Poisson process on $[0, \infty) \times \{0, 1\}^{\infty}$ with intensity measure $dt \otimes \int_{[0,1]} \bigotimes_{n \in \mathbb{N}} P^{(x)} x^{-2} \Lambda(dx)$, where $P^{(x)}$ is the Bernoulli distribution with success probability $x \in (0, 1]$.

When constructing simple Λ -coalescents, even the process \mathcal{P} itself has almost surely finitely many points $(T_j, (K_i^{(j)})_{i \in \mathbb{N}})$ on any set $[0, t] \times \{0, 1\}^{\infty}$ (which we can again order in the first coordinate). As described in [32, Example 19] and analogously to Lemma 1, we can construct each (potential) merger at point $(T_j, (K_i^{(j)})_{j \in \mathbb{N}})$ of a simple Λ -coalescent as follows (while between jumps, we wait independent $\exp(\mu_{-2})$ times). First choose $P_i \in (0, 1]$ from $\mu_{-2}^{-1}x^{-2}\Lambda(dx)$, we have $E(P_i) =$ $\mu_{-2}^{-1}\int_{[0,1]}x^{-1}\Lambda(dx) = \alpha$. Then, throw independent $\cos(K_i^{(j)})_{i \in \mathbb{N}}$ with probability P_i for 'heads' (=1) for each block present and merge all blocks whose coins came up 'heads'. Again, $(P_i)_{i \in \mathbb{N}}$ are i.i.d. and the 'coins' $K_i^{(j)}$ are exchangeable (P_i) indicators. Analogously to above, we thus have

Lemma 2. Let Λ be a finite measure on [0, 1] fulfilling (3). For the Poisson process $\mathcal{P} = (T_j, (K_i^{(j)})_{i \in \mathbb{N}})_{j \in \mathbb{N}}, ((K_i^{(j)})_{i \in \mathbb{N}})_{j \in \mathbb{N}}$ is an i.i.d. sequence (in j) of sequences of exchangeable(P_j) indicators, where $(P_j)_{j \in \mathbb{N}}$ are i.i.d. with $P_1 \stackrel{d}{=} \mu_{-2}^{-1} x^{-2} \Lambda(dx)$. In particular, $E(P_i) = \alpha$.

Since many proofs will build on the properties of different sets of exchangeable indicators, we collect some well-known properties in the following

Lemma 3. Let $(K_i)_{i \in \mathbb{N}}$ be exchangeable(X) indicators.

- a) We have $\lim_{n\to\infty} \frac{1}{n} \sum_{i=1}^{n} K_i = X$ almost surely. X is almost surely unique.
- b) If $(K_i)_{i \in \mathbb{N}}$ is independent of a σ -field \mathcal{F} , X is, too.

c) Let $(L_i)_{i \in \mathbb{N}}$ be exchangeable(Y) indicators, independent of $(K_i)_{i \in \mathbb{N}}$. Then, $(K_iL_i)_{i \in \mathbb{N}}$ are exchangeable(XY) indicators and X, Y are independent.

Proof. These properties essentially follow from the de Finetti representation of an infinite series of exchangeable variables as conditionally i.i.d. variables. The lemma is a collection of well-known properties as e.g. described in [3, Sections 2 and 3], arguments of which we use in the following.

An infinite exchangeable sequence is conditionally i.i.d. given an almost surely unique random measure α . This measure is the weak limit of the empirical measures, in our case, $n^{-1} \sum_{i=1}^{n} \delta_{K_i}$, which has limit $X'\delta_1 + (1-X')\delta_0$ for some random variable X' with values in [0, 1]. Given α , the indicators are α -distributed. However, since Xgives the success probability of each Bernoulli coin, we have X = X' almost surely, so X is almost surely unique. The rest of a) is just the strong law of large numbers e.g. from [3, 2.24] ($E(K_1) \le 1$), the limit is X'. Part b) follows from measure theory since the limit is measurable in the σ -field spanned by the summed variables. For c), we again check Pitman's condition [32, Eq. 27]. Let $M \subseteq [m]$. We have that X, Y are independent from b). With $P(K_i = L_i = 1|X, Y) = XY$ almost surely,

$$P(\{i \in [m] : K_i L_i = 1\} = M) = E(P(\{i \in [m] : K_i L_i = 1\} = M | X, Y))$$
$$= E((XY)^{|M|}(1 - XY)^{m - |M|}),$$

since given X, Y, both $(K_i)_{i \in \mathbb{N}}$ and $(L_i)_{i \in \mathbb{N}}$ are independent. This shows c).

3 When does a *Ξ*-coalescent with dust stay infinite?

A crucial assumption for our results is that the Ξ -coalescent Π has almost surely infinitely many blocks that may merge in the mergers where 1 participates in. The property

 $P(\Pi_t \text{ has infinitely many blocks } \forall t > 0) = 1$

is called staying infinite, while $P(\Pi_t \text{ has finitely many blocks } \forall t > 0) = 1$ is the property of coming down from infinity. These properties have been thoroughly discussed for Ξ -coalescents, see e.g. [33, 27] and [20].

We recall the condition for Ξ -coalescents with dust to stay infinite.

Lemma 4. Let $\Delta_f := \{x \in \Delta : x_1 + \dots + x_k = 1 \text{ for some } k \in \mathbb{N}\}$ and Ξ be a finite measure on Δ fulfilling Eq. (2). The Ξ -coalescent stays infinite if and only if $\Xi(\Delta_f) = 0$. If $\Xi(\Delta_f) > 0$, then the Ξ -coalescent has infinitely many blocks until the first jump of f_1 almost surely and the Ξ -coalescent neither comes down from infinity nor stays infinite.

Proof. Let $\Delta^* := \{x \in \Delta : |x| = 1\}$. All Ξ -coalescents considered are constructed via the Poisson construction with Poisson point process \mathcal{P} .

First, assume $\mathcal{Z}(\Delta^*) = 0$. We recall the (well-known) property that for a \mathcal{Z} coalescent with dust $\mathcal{Z}(\Delta^*) = 0$ is equivalent to $P(S_t > 0 \forall t) = 1$, where S_t is
the asymptotic frequency of the dust component. We use the remark on [12, p.1091]:
For \mathcal{Z} -coalescents with dust, $(-\log S_t)_{t\geq 0}$ is a subordinator. The subordinator jumps
to ∞ (corresponds to $S_t = 0$) if and only if for its Laplace exponent Φ , we have

 $\lim_{\eta \searrow 0} \Phi(\eta) > 0$. For a Ξ -coalescent with dust we have $\lim_{\eta \searrow 0} \Phi(\eta) = \int_{\Delta^*} \nu_0(dx)$. Hence, $\Xi(\Delta^*) = 0$ almost surely guarantees infinitely many singleton blocks for all $t \ge 0$, so the corresponding Ξ coalescent stays infinite.

Now assume $\Xi(\Delta^*) > 0$. The subordinator $(-\log S_t)_{t\geq 0}$ jumps from finite values $(S_t > 0)$ to ∞ $(S_t = 0)$ after an exponential time with rate $\nu_0(\Delta^*)$. This shows that the Ξ -coalescent does not come down from infinity. Assume further that $\Xi(\Delta_f) = 0$. Then, [33, Lemma 31] shows that the Ξ -coalescent either comes down from infinity or stays infinite, so it stays infinite.

Finally, assume $\mathcal{Z}(\Delta_f) > 0$. Split \mathcal{P} into independent Poisson processes $\mathcal{P}'_1 :=$ $\{(T, (K_i)_{i \in \mathbb{N}}) \in \mathcal{P} : \kappa \in \Delta_f\}$ and $\mathcal{P}'_2 := \{(T, (K_i)_{i \in \mathbb{N}}) \in \mathcal{P} : \kappa \notin \Delta_f\}$, where $\kappa := (\lim_{n \to \infty} n^{-1} \sum_{i \in [n]} 1_{\{K_i = j\}})_{i \in \mathbb{N}}$ (again restriction theorem [25, p.17], Lemma 3 shows κ exists almost surely). Their intensity measures are defined as in Eq. (11), but using $v'_1(\cdot) := v_0(\cdot \cap \Delta_f)$ and $v'_2 := v_0 - v'_1$ instead of v_0 . Since $v'_1(\Delta_f) \leq v'_1(\Delta_f)$ $\mu_{-1} < \infty$, for any t > 0 there are almost surely finitely many $P \in \mathcal{P}'_1$ with T < t. Consider such $P = (T, (K_i)_{i \in \mathbb{N}})$ with T smallest. Observe that until T, we can construct the Ξ -coalescent using only the points of \mathcal{P}'_2 , which is the construction of a Ξ' -coalescent with $\Xi'(dx) := (x, x)\nu'_2(dx)$. Since $\int_{\Delta} |x|\nu'_2(dx) < \mu_{-1} < \infty$ and $\Xi'(\Delta_f) = 0$, the proof steps above show that the Ξ -coalescent has infinitely many blocks until T. Now consider the merger at time T. The form of v'_1 ensures that $(K_i)_{i \in \mathbb{N}}$ can only take finitely many values, and Lemma 3a) ensures that infinitely many K_i 's show each value. Thus, all blocks present before time T are merged at T into a finite number of blocks (given by which K_i 's show the same number). This shows that if $\mathcal{Z}(\Delta_f) > 0$, the Z-coalescent stays neither infinite nor comes down from infinity. Additionally, this shows that either the block of 1 already merged at least once before T or it merges at T, thus there are infinitely many blocks before the first merger of 1.

4 The block of 1 in Ξ-coalescents with dust – proofs and remarks

Proof of Theorem 1. As in Lemma 1, split the Poisson point process \mathcal{P} used to construct the Ξ -coalescent in \mathcal{P}_1 and \mathcal{P}_2 . We also use the notation from Lemma 1 and its proof. The block of 1 in the Ξ -*n*-coalescent for any $n \in \mathbb{N}$ can only merge at times *t* for which there exists a Poisson point $(T, (K_i)_{i \in \mathbb{N}}) \in \mathcal{P}_1$. Lemma 1 states that the set of times *T* forms a homogeneous Poisson process with rate μ_{-1} . This shows that potential jump times are separated by countably many independent $\text{Exp}(\mu_{-1})$ random variables. Kingman's correspondence yields that f_1 exists almost surely at each potential jump time. To see this, observe that even though the partition of \mathbb{N} induced by the Poisson construction is not exchangeable, the partition on $\mathbb{N} \setminus \{1\}$ is, and the asymptotic frequencies of the former and the latter coincide. Since f_1 is by definition constant between these jump points, f_1 has càdlàg paths almost surely. Since any blocks change by mergers, f_1 is increasing.

The value of f_1 at 0 follows by definition. Since Π stays infinite (see Lemma 4), at each $P \in \mathcal{P}_1$ infinitely many blocks can potentially merge. Lemma 1 shows that the indicators of whether blocks present immediately before P merge with the block of 1 are exchangeable(X) indicators with X > 0 almost surely. Then, Lemma 3 ensures that a positive fraction of them almost surely does, causing f_1 to jump (since

a positive fraction of merging blocks has positive frequency). Thus, every Poisson point leads to a merger almost surely, which shows that f_1 jumps at all potential jump times described above. Since, for all t, either $S_t > 0$ or non-dust blocks not including 1 exist (having asymptotical frequency > 0), $f_1(t) < 1$ for all $t \ge 0$.

We consider the jump chain of f_1 . Set $X_1 := f_1[1]$. Since $f_1[k] \in (0, 1)$ for all $k \in \mathbb{N}$ and f_1 increases, $f_1[k+1] = f_1[k] + (1 - f_1[k])X_{k+1}$ for $X_{k+1} \in (0, 1)$. Iterating this yields $f_1[k] = \sum_{i=1}^k X_i \prod_{j=1}^{i-1} (1 - X_j)$ for $k \ge 2$. The properties of $(X_k)_{k\in\mathbb{N}}$ follow from the Poisson construction and Lemma 1. Consider the blocks present at time T_k -, where the *k*th Poisson point of \mathcal{P}_1 is $P_k = (T_k, (K_i^{(k)})_{i\in\mathbb{N}})$. The block with least element *i* merges with the block of 1 if $K_i^{(k)} = K_1^{(k)}$. Consider the *k*th Poisson point at time T_k . X_k gives the fraction of the asymptotic frequency of non-singleton blocks and singleton blocks at time T_k -, i.e. the fraction of $1 - f_1(T_{k-})$, that is merged with the block of 1 at T_k . Denote $L_i^{(k-)} := 1_{\{i\} \text{ is a block at } T_k-\}$. Then, recording the asymptotic frequencies of merged non-singleton and singleton blocks,

$$X_{k} = \frac{1}{1 - f_{1}(T_{k} - i)} \left(\sum_{i \ge 2} \mathbb{1}_{\{K_{1}^{(k)} = K_{i}^{(k)}\}} f_{i}(T_{k} - i) + \lim_{n \to \infty} \frac{1}{n} \sum_{i=2}^{n} \mathbb{1}_{\{K_{1}^{(k)} = K_{i}^{(k)}\}} L_{i}^{(k-i)} \right).$$

Since by construction, $\Pi_{T_k-} \setminus \{1\}$ is an exchangeable partition of $\mathbb{N} \setminus \{1\}$, $(L_i^{(k-)})_{i \in \mathbb{N}}$ are exchangeable (S_{t-}) indicators with $S_{t-} = 1 - \sum_{i=1}^{\infty} f_i(T_k-)$. Recall that Lemma 1 tells us that $(1_{\{K_1^{(k)} = K_i^{(k)}\}})_{i \geq 2}$ are exchangeable(X') indicators with $X' \stackrel{d}{=} Q$. $(K_i^{(k)})_{i \in \mathbb{N}}$ is independent from $(\Pi_t)_{t < T_k}$, since the Poisson points of \mathcal{P}_1 are i.i.d., so Lemma 3 c) and a) show

$$X_k \stackrel{a.s.}{=} \sum_{i \ge 2} \mathbb{1}_{\{K_1^{(k)} = K_i^{(k)}\}} \frac{f_i(T_k)}{1 - f_1(T_k)} + X' \frac{1 - \sum_{i=1}^{\infty} f_i(T_k)}{1 - f_1(T_k)}.$$
 (14)

The independence of $(K_i^{(k)})_{i \in \mathbb{N}}$ from $(\Pi_t)_{t < T_k}$ is also crucial for the next two equations. Compute, with $P(K_1^{(k)} = K_i^{(k)}) = E(X') = \gamma$ for $i \in \mathbb{N}$,

$$E(X_k) = \sum_{i\geq 2} P\left(K_1^{(k)} = K_i^{(k)}\right) E\left(\frac{f_i(T_k)}{1 - f_1(T_k)}\right) + E\left(X'\right) E\left(\frac{1 - \sum_{i=1}^{\infty} f_i(T_k)}{1 - f_1(T_k)}\right)$$
$$= \gamma E\left(\frac{1 - f_1(T_k)}{1 - f_1(T_k)}\right) = \gamma.$$

Analogously, for $l < k, X_l$ only depends on Poisson points P_1, \ldots, P_l , so

$$E(X_k X_l) = E\left(\sum_{i\geq 2} 1_{\{K_1^{(k)} = K_i^{(k)}\}} \frac{f_i(T_k)}{1 - f_1(T_k)} X_l + X' \frac{1 - \sum_{i=1}^{\infty} f_i(T_k)}{1 - f_1(T_k)} X_l\right)$$
$$= \gamma E\left(\frac{1 - f_1(T_k)}{1 - f_1(T_k)} X_l\right) = \gamma^2,$$

showing that X_k, X_l are uncorrelated. An analogous computation shows that $E(\prod_{i \in \{l_1, \dots, l_m\}} X_{l_i}) = \prod_{i \in \{l_1, \dots, l_m\}} E(X_{l_i})$ for distinct $l_1, \dots, l_m \in \mathbb{N}$. With this,

$$E(f_1[k]) = \sum_{i=1}^{k} E(X_i) \prod_{j=1}^{i-1} (1 - E(X_j)) = \sum_{i=1}^{k} \gamma (1 - \gamma)^{i-1} = 1 - (1 - \gamma)^k$$

To prove $\lim_{t\to\infty} f_1(t) = 1$ almost surely, observe that f_1 is bounded and increasing, thus $\lim_{t\to\infty} f_1(t)$ exists. Monotone convergence and $\lim_{t\to\infty} E(f_1(t)) = \lim_{k\to\infty} E(f_1[k]) = 1$ show the desired. Note that $(X_k)_{k\in\mathbb{N}}$ is in general neither independent nor identically distributed, see Section 6.

Proof of Corollary 1. By the Poisson construction the block of 1 for $\Pi^{(n)}$ can only merge at times given by Poisson points in \mathcal{P}_1 . Consider $(T_1, (K_i^{(1)})_{i \in \mathbb{N}}) \in \mathcal{P}_1$. While T_1 is the time of the first jump of f_1 (see the proof of Theorem 1), there does not necessarily need to be a merger of {1} in the *n*-coalescent $\Pi^{(n)}$, if we have $K_1^{(1)} \neq K_i^{(1)}$ for the least elements *i* of all other blocks of $\Pi^{(n)}$ immediately before T_1 . However, Lemma 1 shows that $(1_{\{K_1^{(1)}=K_i^{(1)}\}})_{i\geq 2}$ are exchangeable indicators. The mean $n^{-1}\sum_{i=2}^n 1_{\{K_1^{(1)}=K_i^{(1)}\}}$, as argued in the proof of Theorem 1, converges to an almost surely positive random variable for $n \to \infty$. As shown in Lemma 4, any Ξ -coalescent with dust has infinitely many blocks almost surely before T_1 . Thus, there exists N, a random variable on \mathbb{N} , so that 1 is also merging at time T_1 in $\Pi^{(n)}$ for $n \ge N$ almost surely. This yields $\lim_{n\to\infty} n^{-1}M_n = \lim_{n\to\infty} n^{-1}|B_1(T_1) \cap [n]| = f_1(T_1) = f_1[1]$ almost surely. All further claims follow from Theorem 1.

Remark 5. Let $Q^{(n)}$ be the number of blocks merged at the first collision of the block of 1 in a Λ -n-coalescent with dust. [34, 1.4] shows that $n^{-1}Q^{(n)}$ converges in distribution. We argue that this convergence also holds in L^p for all p > 0 and, for simple Λ -n-coalescents, almost surely.

The proof of Corollary 1 shows that $(T_1, (K_i^{(1)})_{i \in \mathbb{N}}) \in \mathcal{P}_1$ causes the first merger in the n-coalescent for n large enough (almost surely, but since $n^{-1}Q^{(n)} \in [0, 1]$ for all n, convergence in L^p is not affected by the null set excluded). Split $Q^{(n)}$ into $Q_0^{(n)}$, the number of non-singleton blocks and $Q_1^{(n)}$, the number of singleton blocks merged at T_1 . For the limit, we can ignore the non-singleton blocks merged. To see this, recall $Q_0^{(n)} \leq K_n$, where K_n is the total number of mergers for the Λ -n-coalescent, since a non-singleton block has to be the result of a merger. [12, Lemma 4.1] tells us that $n^{-1}K_n \to 0$ in L^1 for $n \to \infty$ for Ξ -coalescents with dust. This shows that the L^1 limit of $n^{-1}Q^{(n)}$ is the same as of the one of $n^{-1}Q_1^{(n)}$. $n^{-1}Q_1^{(n)}$ already appeared in the part of the proof of Theorem 1 leading to Eq. (14), its limit almost surely exists and equals $X' \frac{1-\sum_{i=1}^{\infty} f_i(T_1-)}{1-f_1(T_1-)}$. Since $n^{-1}Q_1^{(n)}$ is bounded in [0, 1] it also converges in L^p , p > 0. For simple Ξ -n-coalescents, [11, Lemma 4.2] shows $n^{-1}K_n \to 0$ almost surely, so in this case the steps above ensure also almost sure convergence of $n^{-1}Q^{(n)}$.

5 The block of 1 in simple Λ-coalescents – proofs and remarks

Proof of Proposition 3. Let $\mathcal{P} := (P_i)_{i \in \mathbb{N}}$ be the coin probabilities coming from the Poisson process used to construct the simple Λ -coalescent Π as described in Section 2. As shown in the proof of Theorem 1, the Poisson point belonging to P_C where 1 first throws 'heads' in the Poisson construction is the Poisson point where f_1 jumps for the first time. We have $P(C = k | \mathcal{P}) = P_k \prod_{i=1}^{k-1} (1 - P_i)$. Integrating the condition and using the independence of $(P_i)_{i \in \mathbb{N}}$ as well as $E(P_1) = \alpha$ (see Lemma 2), we see that *C* is geometrically distributed with parameter α .

To describe $f_1[1]$ at the *C*th merger (Poisson point), recall that the restriction Π_{-1} of Π to $\mathbb{N} \setminus \{1\}$ has the same asymptotic frequencies as Π . Thus, we can see $f_1[1]$ as the asymptotic frequency of the newly formed block of Π_{-1} at the time of the Poisson point P_C . This follows since Π_{-1} has infinitely many blocks before (see Lemma 4) and then, as in the proof of Theorem 1, there will be a newly formed block of Π_{-1} at the *C*th Poisson point (and the unrestricted block in Π includes 1).

We consider Π_{-1} at the *k*th Poisson point with coin probability P_k . For $\{i\} \in \mathbb{N} \setminus \{1\}$ to remain a (singleton) block and not be merged for the first k - 1 mergers and then to be merged at the *k*th, we need $\prod_{j \in [k-1]} (1 - K_i^{(j)}) = 1$ and $K_i^{(k)} = 1$. $(1_{\{\prod_{j \in [k-1]} (1 - K_i^{(j)}) = 1, K_i^{(k)} = 1\}})_{i \in \mathbb{N}}$ are exchangeable $(P_k \prod_{j \in [k-1]} (1 - P_j))$ indicators. Let

$$S_{k} = \left\{ i \in \mathbb{N} \setminus \{1\} : \prod_{j \in [k-1]} \left(1 - K_{i}^{(j)}\right) = 1, K_{i}^{(k)} = 1 \right\}$$

be the set of $i \in \mathbb{N} \setminus \{1\}$ whose first merger is the *k*th overall merger. We call S_k the *k*th singleton set (of Π_{-1}). From the strong law of large numbers for exchangeable indicators, see Lemma 3a), we directly have that S_k has asymptotic frequency $P_k \prod_{i \in [k-1]} (1 - P_i)$ almost surely.

Now, consider the asymptotic frequency $f^*[k]$ of the newly formed block at the *k*th merger of Π_{-1} . By construction, there is only one newly formed block at each merger. S_k is a part of the newly formed block. Any other present block with more than two elements (non-singleton block) is merged if and only if its indicator $K_i^{(k)} = 1$ (we order by least elements). For k = 1, the newly formed block is S_1 . For k = 2, it is either S_2 or $S_1 \cup S_2$, if the coin of the the block S_1 formed in the first merger comes up 'heads'.

Applied successively, this shows that the newly formed block at the *k*th merger consists of a union of a subset of the singleton sets $(S_{k'})_{k' < k}$ and the set S_k . For its asymptotic frequency, we have

$$f^*[k] = \sum_{i=1}^k B_i^{(k)} P_i \prod_{j \in [i-1]} (1 - P_j) > 0,$$
(15)

where the $B_i^{(k)}$, $i \in [k]$, are non-independent Bernoulli variables which are 1 if the *i*th singleton set S_i is a part of the newly formed block at the *k*th merger of Π_{-1} .

If $\Lambda(\{1\}) > 0$, $P_k = 1$ is possible. In this case, at the *k*th Poisson point all remaining singletons form S_k and all blocks present at merger k - 1 merge with S_k .

There are no mergers at Poisson points P_l , l > k, so we do not consider Eq. (15) for l > k.

We have $f_1[1] = f^*[C]$. Given \mathcal{P} , $(f^*[k])_{k \in \mathbb{N}}$ is independent of C. Thus, Eq. (9) is implied by Eq. (15).

Assume $\Lambda(\{1\}) = 0$. For $(B_i^{(k)})_{k \in \mathbb{N}, i \in [k]}$, we have $B_k^{(k)} = 1$ for all $k \in \mathbb{N}$ since the *k*th singleton set is formed at the *k*th Poisson point and is a part of the newly formed block. The coins thrown at the *k*th Poisson point to decide whether other singleton sets S_i , S_j with i, j < k are also parts of the newly formed block are either independent given \mathcal{P} when they are in different blocks, or identical when they are in the same block. The set S_i uses the coin of the block newly formed at the *i*th merger. Let I(i) be the Poisson point at which this block merges again (and S_i with it). At the I(i)th Poisson point and for all further Poisson points indexed with $j \ge I(i)$, we have $B_i^{(j)} = B_{I(i)}^{(j)}$, since the singleton sets S_i and $S_{I(i)}$ are in the same block for mergers $j \ge I(i)$.

The property (i) of I(i) in the proposition follow directly from its definition as the minimum number of coin tosses until the first comes up 'heads'. The property (ii) is just integrating (i) and using that $(P_i)_{i \in \mathbb{N}}$ are i.i.d. with $E(P_1) = \alpha$ (see Lemma 2), the conditional independence is the conditional independence of coin tosses of distinct blocks from the Poisson construction. To see Eq. (10), observe that S_i for i < j is a part of the newly formed block at the *j*th merger of the Λ -coalescent $(i \in J)$ if and only if $I(i) \in J$. If $I(i) \in J$, either we have I(i) = j, so S_i is merged for the first time after it has been formed at the *j*th merger, or we have that I(i) < jwhich means that it has already merged with at least one other singleton set and that, as parts of the same block, they both again merged at the *j*th merger. If $I(i) \notin J$, the singleton set S_i neither merges at the *j*th merger for the first time after being formed nor merges with any other singleton set before that is then merging at the *j*th merger, so S_i is not a part of the newly merged block at the *j*th merger.

If $\Lambda(\{1\}) = 0$, the arguments hold true for all $j \in \mathbb{N}$. If $\Lambda(\{1\}) > 0$ this holds true for all $i, j \leq K := \min_{k \in \mathbb{N}} \{P_k = 1\} (< \infty \text{ almost surely})$, where all singleton sets merge and $f^*[K] = 1$. However, in this case $C \leq K$, so we still can establish Eq. (9).

Remarks 6.

- (I(i))_{i∈ℕ} is useful to construct the asymptotic frequencies of the A-coalescent. Given P, at the kth merger, there are the singleton sets (S_j)_{j∈[k]} with almost sure frequencies P_j ∏_{i∈[j-1]}(1 − P_i) which were already formed in the k collisions, and unmerged singleton blocks with frequency ∏_{i∈[k]}(1 − P_i). Using (I(i))_{i∈[k]}, we can indicate which singleton sets form a block. S_i is a single block if I(i) > k, if I(i) ≤ k it is a part of a block where S_{I(i)} is also a part of. This can be seen as a discrete version of the construction of the A-coalescent from the process of singletons as described in [15, Section 6.1]
- The variables $(I(i))_{i \in \mathbb{N}}$ are useful to express other quantities of the Λ -coalescent. For instance, the number of non-singleton blocks in a simple Λ -coalescent at the kth merger is given by $k \sum_{i \in [k-1]} 1_{\{I(i) \leq k\}}$.

To prove Proposition 2, we need the following result.

Lemma 5. For $p \in [\frac{1}{2}, 1)$, each $x \in \mathcal{M}_p$ from Eq. (6) has a unique representation in \mathcal{M}_p .

Proof. We adjust the proof of [4, Theorem 7.11]. Assume that $x \in \mathcal{M}_p$ has two representations $x = \sum_{i \in \mathbb{N}} b_i pq^{i-1} = \sum_{i \in \mathbb{N}} b'_i pq^{i-1}$ with $b_i \neq b'_i$ for at least one *i*. Let i_0 be the smallest integer with $b_{i_0} \neq b'_{i_0}$. Without restriction, assume $b_{i_0} - b'_{i_0} = 1$. Then,

$$0 = \sum_{i \in \mathbb{N}} b_i p q^{i-1} - \sum_{i \in \mathbb{N}} b'_i p q^{i-1} = p q^{i_0 - 1} + \sum_{i > i_0} (b_i - b'_i) p q^{i-1}.$$

Thus, $pq^{i_0-1} = \sum_{i>i_0} (b'_i - b_i) pq^{i-1} < \sum_{i>i_0} pq^{i-1} = pq^{i_0}$, simplifying to p < q, in contradiction to the assumption $p \ge \frac{1}{2}$.

Proof of Proposition 2. From Eq. (15) we see that f_1 only takes values in \mathcal{M}_p , since $P_k = p$ for all $k \in \mathbb{N}$ and $C < \infty$ almost surely. Recall the definition of the singleton sets S_i and their properties from the proof of Proposition 3. The asymptotic frequency of S_i is pq^{i-1} almost surely. Lemma 5 ensures that there is a unique representation $f_1[l] = x = \sum_{i=1}^{\infty} b_i pq^{i-1}$ in \mathcal{M}_p , let $J := \{i \in \mathbb{N} : b_i = 1\}$ and $j := \max J$. This means that $f_1[l] = x$ is equivalent to that the block of 1 at its *l*th jump consists of the union of all S_i with $i \in J$ and 1. This also shows that the *l*th jump of f_1 is at the *j*th jump of the Dirac coalescent, since if f_1 jumps at the *k*th merger of the Dirac coalescent, the newly formed block includes S_k .

Since $P_i = p$ for all $i \in \mathbb{N}$, we have $\alpha = p$ and Eq. (9) simplifies to $f_1[1] = \sum_{i=1}^{C} B_i^{(C)} pq^{i-1}$, where $C \stackrel{d}{=} Geo(p)$ is independent from $(B_i^{(k)})_{k \in \mathbb{N}, i \in [k]}$. The latter fulfil

$$P(B_i^{(j)} = b_i \forall i \in [j-1]) = \prod_{i \in J \setminus \{j\}} P(Y+i \in J) \prod_{i \in [j] \setminus J} P(Y+i \notin J) \quad (16)$$

with $Y \stackrel{d}{=} Geo(p)$, since the joint distribution in Proposition 3 again simplifies, we can ignore the conditioning and $I(i) - i \stackrel{d}{=} Geo(p)$ for all $i \in \mathbb{N}$.

Since $f_1[1] = x$ uniquely determines the values of C and $(B_i^{(C)})_{i \in [C]}$, we have

$$P(f_1[1] = x) = P(C = j)P(B_1^{(j)} = b_1, \dots, B_{j-1}^{(j)} = b_{j-1}),$$
(17)

which shows Eq. (7) when we insert the distributions expressed in terms of their geometric distributions.

In order to verify that the jump chain $(f_1[i])_{i \in \mathbb{N}}$ is Markovian, we show that $f_1[1], \ldots, f_1[l]$ does not contain more information on $f_1[l + 1]$ than $f_1[l]$ does. Without restriction, assume that the *l*th jump $f_1[l]$ of f_1 takes place at the *k*th jump of the Dirac coalescent. Then, $f_1[l+1]$ is constructed from the blocks present after the *k*th merger. For each subsequent Poisson point P_{k+1}, \ldots , blocks present are merged if their respective coins come up 'heads' until (and including), at $P_{k'}$, the coin of the block of 1 comes up 'heads' for the first time since P_k . Thus, only information about the block partition at merger *k* can change the law of the next jump. $f_1[l] = x$ gives the information which singleton sets S_1, \ldots, S_k are parts of the block of 1 at merger

k of Π and which are not. $f_1[l] = x$ contains no information about how the other singleton sets, S_i with $b_i = 0$, are merged into blocks at collisions before *k* apart from that it tells us that $B_i^{(j)} = 0$ for $j \in J$ and $i \notin J$, which means that all S_i with $i \notin J$ did not merge at the *j*th collisions, $j \in J$. This is due to that any S_i with $B_i^{(j)} = 1$ would merge with the newly formed block at merger *j* and thus would be in a block with S_j and also in the block of 1 at merger *k*. However, analogously we see that knowing $f_1[1], \ldots, f_1[l]$ does not give any additional information about the block structure at the *k*th merger, but only how the set of S_i which are in the block of 1 at merger *k* behaved at the earlier mergers *J*. Thus, $(f_1[l])_{l\in\mathbb{N}}$ is Markovian. However, $(f_1(t))_{t>0}$ is not Markovian. In order to see this consider, for $0 < t_0 < t_1 < t_2$,

$$p(t_2, t_1, t_0) := P(f_1(t_2) = p + pq^2 | f_1(t_1) = p, f_1(t_0) = 0)$$

=
$$\frac{P(f_1(t_2) = p + pq^2, f_1(t_1) = p, f_1(t_0) = 0)}{P(f_1(t_1) = p, f_1(t_0) = 0)}.$$

We will show that $p(t_2, t_1, t_0)$ depends on t_0 , which shows that f_1 is not Markovian.

We can express all events in terms of the independent waiting times for Poisson points, i.e. the successive differences between the first component *T* of the Poisson points $(T, (K_i)_{i \in \mathbb{N}}) \in \mathcal{P}$. Here, we use the split of the Poisson points into the independent Poisson point processes \mathcal{P}_1 and \mathcal{P}_2 from Lemma 1. The waiting times between points in \mathcal{P}_1 are $\exp(\mu_{-1})$ -distributed, the waiting times between points in \mathcal{P}_2 are $\exp(\mu_{-2} - \mu_{-1})$ -distributed, see Lemma 1 and Remark 4. We will relabel $\tau = \mu_{-1}$ and $\rho = \mu_{-2} - \mu_{-1}$ for a clearer type face. Let T_1, T_2, \ldots be the waiting times between points in \mathcal{P}_1 and T'_1, T'_2, \ldots be the waiting times between points in \mathcal{P}_2 . All waiting times are independent one from another. We recall that for $T \stackrel{d}{=} \exp(\alpha)$, $P(T > a) = e^{-\alpha a}$ and $P(T \in (a, a + b]) = e^{-\alpha a}(1 - e^{-\alpha b})$ for $a, b \ge 0$.

The event $\{f_1(t_1) = p, f_1(t_0) = 0\}$ means that the first jump of f_1 adds the singleton set S_1 at a time in $(t_0, t_1]$. Thus, there has to be only a single point of \mathcal{P}_1 with first component $T_1 \leq t_1$ and the smallest time T'_1 of points of \mathcal{P}_2 has to be greater than T_1 . We compute, conditioning on T_1 for the third equation,

$$P(f_{1}(t_{1}) = p, f_{1}(t_{0}) = 0) = P(t_{0} < T_{1} \le t_{1} < T_{1} + T_{2}, T_{1} < T_{1}')$$

$$= \int_{t_{0}}^{t_{1}} P(T_{2} > t_{1} - x) P(T_{1}' > x) \tau e^{-\tau x} dx$$

$$= \int_{t_{0}}^{t_{1}} e^{-\tau(t_{1} - x)} e^{-\rho x} \tau e^{-\tau x} dx$$

$$= \frac{\tau}{\rho} e^{-\tau t_{1}} \int_{t_{0}}^{t_{1}} \rho e^{-\rho x} dx = \frac{\tau}{\rho} e^{-\tau t_{1}} (e^{-\rho t_{0}} - e^{-\rho t_{1}})$$

Analogously, we compute (by conditioning on T_1 , T_2 for the second equality)

$$P(f_1(t_2) = p + pq^2, f_1(t_1) = p, f_1(t_0) = 0)$$

= $P(t_0 < T_1 \le t_1 < T_1 + T_2 \le t_2 < T_1 + T_2 + T_3, T_1 < T_1' \le T_1 + T_2 < T_2')$
= $\int_{t_0}^{t_1} \int_{t_1-x}^{t_2-x} P(T_3 > t_2 - x - y, T_1' \in (x, x+y], T_2' > x+y)\tau^2 e^{-\tau x} e^{-\tau y} dy dx$

$$= \tau^2 \int_{t_0}^{t_1} \int_{t_1-x}^{t_2-x} e^{-\tau(t_2-x-y)} e^{-\rho x} (1-e^{-\rho y}) e^{-\rho(x+y)} e^{-\tau x} e^{-\tau y} dy dx$$

= $\frac{\tau^2}{\rho} e^{-\tau t_2} \bigg[\rho^{-1} (e^{-\rho t_1} - e^{-\rho t_2}) (e^{-\rho t_0} - e^{-\rho t_1}) \bigg] - \frac{1}{2} (e^{-2\rho t_1} - e^{-2\rho t_2}) (t_1 - t_0) \bigg].$

Taking the ratio shows that

$$p(t_2, t_1, t_0) = \frac{\tau}{\rho} e^{-\tau(t_2 - t_1)} \left(e^{-\rho t_1} - e^{-\rho t_2} \right) - \underbrace{\frac{\tau}{2} e^{-\tau(t_2 - t_1)} \frac{e^{-2\rho t_1} - e^{-2\rho t_2}}{e^{-\rho t_0} - e^{-\rho t_1}}}_{\neq 0} (t_1 - t_0)$$

depends on t_0 , so f_1 is not Markovian.

Remark 7. Our proof of Proposition 2 relies on the unique representation in \mathcal{M}_p . This means that it also holds true for all $p \in (0, 2^{-1})$ where each $x \in \mathcal{M}_p$ has a unique representation in \mathcal{M}_p , e.g. for all transcendental p. If the representation is not unique, Eq. (16) is still correct, but the right side of Eq. (17) does not show $P(f_1[1] = x)$. Instead, the latter shows the contribution to $P(f_1[1] = x)$ from the paths of f_1 which fulfil C = j, $B_1^{(j)} = b_1, \ldots, B_{j-1}^{(j)} = b_{j-1}$ (recall that j, b_1, \ldots, b_{j-1} depend on the representation of x). Moreover, $P(f_1[1] = x)$ then is the sum over $P(C = j)P(B_1^{(j)} = b_1, \ldots, B_{j-1}^{(j)} = b_{j-1})$ for the tuples j, b_1, \ldots, b_{j-1} coming from the different representations of x (the sets of paths are disjoint if the parameter sets $(j, b_1, \ldots, b_{j-1})$ differ). Since the proof of our results on the Markov property of both f_1 and its jump chain also rely on the unique representation of x (to read off which blocks merged when), the proof does not extend if p does not allow a unique representation of x.

6 Example

We provide a concrete example showing that the random variables $(X_k)_{k \in \mathbb{N}}$ from Theorem 1 are, in general, neither independent nor identically distributed.

Choose $\Lambda = \delta_{\frac{1}{2}}$ and consider f_1 in the corresponding Λ -coalescent. Recall that $f_1[l] = x \in \mathcal{M}_{\frac{1}{2}}$ already fixes which singleton sets S_k are parts of the block of 1 at its *l*th merger and which are not. First, assume $f_1[1] = X_1 = \frac{5}{8} = \frac{1}{2} + \frac{1}{2^3} \in \mathcal{M}_{\frac{1}{2}}$, which means that the coin of 1 comes up 'heads' for the first time at the third Poisson point and the block of 1 is $S_1 \cup S_3$, while S_2 is a block of its own (an event happening with probability > 0). Assume further $f_1[2] = \frac{11}{16} = \frac{5}{8} + \frac{1}{16}$. This sets $X_2 = (f_1[2] - f_1[1])/(1 - X_1) = \frac{1}{6} \notin \mathcal{M}_{\frac{1}{2}}$. We read off that the coin of the block of 1 also comes up 'heads' at the fourth collision, where the block of 1 merges with S_4 . We also see that the coin of the only other block S_2 comes up 'tails'. We thus have, since we throw fair coins, $P(X_2 = \frac{1}{6}|X_1 = \frac{5}{8}) = P(f_1[2] = \frac{11}{16}|f_1[1] = \frac{5}{8}) = \frac{1}{4}$. Since $X_1 = f_1[1] \in \mathcal{M}_{\frac{1}{2}}$ for any realisation, X_1 and X_2 have different distributions. To see also non-independence, consider $f_1[1] = X_1 = \frac{1}{2}$ (coin of 1 comes up 'heads' at first coin toss, block of 1 is S_1 , occurs with probability $\frac{1}{2}$). In this case $P(X_2 = \frac{1}{6}|X_1 = \frac{1}{2}) = 0$, since $f_1[2] = X_1 + (1 - X_1)X_2 = \frac{7}{12} \notin \mathcal{M}_{\frac{1}{2}}$.

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