# On the size of the block of 1 for $\Xi$-coalescents with dust 

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#### Abstract

We study the frequency process $f_{1}$ of the block of 1 for a $\Xi$-coalescent $\Pi$ with dust. If $\Pi$ stays infinite, $f_{1}$ is a jump-hold process which can be expressed as a sum of broken parts from a stick-breaking procedure with uncorrelated, but in general non-independent, stick lengths with common mean. For Dirac- $\Lambda$-coalescents with $\Lambda=\delta_{p}, p \in\left[\frac{1}{2}, 1\right), f_{1}$ is not Markovian, whereas its jump chain is Markovian. For simple $\Lambda$-coalescents the distribution of $f_{1}$ at its first jump, the asymptotic frequency of the minimal clade of 1 , is expressed via conditionally independent shifted geometric distributions.


Keywords $\Xi$-coalescent, coalescent with dust, Poisson point process, minimal clade, exchangeability
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## 1 Introduction and results

Independently introduced in [33] and [30], $\boldsymbol{\Xi}$-coalescents are exchangeable Markovian processes $\Pi=\left(\Pi_{t}\right)_{t \geq 0}$ on the set of partitions of $\mathbb{N}:=\{1,2, \ldots\}$ whose transitions are due to mergers of partition blocks. The distribution of $\Pi$ is characterised by

[^0]a finite measure $\Xi$ on the infinite simplex
$$
\Delta:=\left\{x=\left(x_{1}, x_{2}, \ldots\right): x_{1} \geq x_{2} \geq \cdots \geq 0,|x| \leq 1\right\}
$$
where $|x|:=\sum_{i \in \mathbb{N}} x_{i}$. We exclude $\Xi=0$, since it leads to a coalescent without coalescence events. $\Xi$-coalescents allow that disjoint subsets of blocks merge into distinct new blocks, hence they are also called coalescents with simultaneous multiple mergers. If $\Xi$ is concentrated on $[0,1] \times\{0\} \times\{0\} \times \cdots$, only a single set of blocks is allowed to merge. Such a coalescent is a $\Lambda$-coalescent, see [32]. In this case, $\Lambda$ is a finite measure on $[0,1]$, the restriction of $\Xi$ on the first coordinate of $\Delta$. The restriction $\Pi^{(n)}$ of $\Pi$ on $[n]:=\{1, \ldots, n\}$ is called the $\Xi-n$-coalescent. Denote the blocks of $\Pi_{t}$ by $\left(B_{i}(t)\right)_{i \in \mathbb{N}}$, where $i$ is the least element of the block (we set $B_{i}(t)=\emptyset$ if $i$ is not a least element of a block). Clearly, $1 \in B_{1}(t)$. We call $B_{1}(t)$ the block of 1 at time $t$. Due to the exchangeability of the $\Xi$-coalescent, Kingman's correspondence ensures that, for every $t \geq 0$, the asymptotic frequencies
\[

$$
\begin{equation*}
f_{i}(t):=\lim _{n \rightarrow \infty} \frac{\left|B_{i}(t) \cap[n]\right|}{n}, \quad i \in \mathbb{N}, \tag{1}
\end{equation*}
$$

\]

exist almost surely, where $|A|$ denotes the cardinality of the set $A$.
The family of $\Xi$-coalescents is a diverse class of processes with very different properties, see e.g. the review [15] for $\Lambda$-coalescents. We will focus on $\Xi$-coalescents with dust, i.e. $\Xi$ fulfils (see [33])

$$
\begin{equation*}
\mu_{-1}:=\int_{\Delta}|x| v_{0}(d x)<\infty \tag{2}
\end{equation*}
$$

where $\nu_{0}(d x)=\Xi(d x) /(x, x)$ with $(x, x):=\sum_{i \in \mathbb{N}} x_{i}^{2}$ for $x=\left(x_{1}, x_{2}, \ldots\right) \in \Delta$. These coalescents are characterised by a non-zero probability that, at any time $t$, there is a positive fraction of $\mathbb{N}$, the dust, that has not yet merged. Note that $i \in \mathbb{N}$ is part of the dust at time $t$ if and only if $\{i\}$ is a block at time $t$, which is called a singleton block. The asymptotic frequency of the dust component is $S_{t}:=1-\sum_{i \in \mathbb{N}} f_{i}(t)$. Having dust is equivalent to $P\left(S_{t}>0\right)>0$ for all $t>0$. We are interested in $\Xi$ coalescents which stay infinite, i.e. which almost surely have an infinite number of blocks for each $t>0$. We will put some further emphasis on simple $\Lambda$-coalescents satisfying

$$
\begin{equation*}
\mu_{-2}:=\int_{[0,1]} x^{-2} \Lambda(d x)<\infty \tag{3}
\end{equation*}
$$

This class includes Dirac coalescents with $\Lambda=\delta_{p}$, the Dirac measure in $p \in(0,1]$. Consider the frequency process $f_{1}:=\left(f_{1}(t)\right)_{t \geq 0}$ of the block of 1 . For $\Lambda$-coalescents, Pitman characterises $f_{1}$ as follows (reproduced from [32], adjusted to our notation).
Proposition 1. [32, Proposition 30] No matter what $\Lambda$, the process $f_{1}$ is an increasing pure jump process with càdlàg paths, $f_{1}(0)=0$ and $\lim _{t \rightarrow \infty} f_{1}(t)=1$. If $\mu_{-1}=\infty$ then almost surely $f_{1}(t)>0$ for all $t>0$ and $\lim _{t \searrow 0} f_{1}(t)=0$. If $\mu_{-1}<\infty$ then $f_{1}$ starts by holding at zero until an exponential time with rate $\mu_{-1}$, when it enters $(0,1]$ by a jump, and proceeds thereafter by a succession of holds and jumps, with holding rates bounded above by $\mu_{-1}$.

Moreover, in [32, Section 3.9], a general formula for the moments of $f_{1}(t)$ for fixed $t>0$ is provided.

For two particular coalescents without dust, further properties of $f_{1}$ are known. For Kingman's $n$-coalescent $\left(\Lambda=\delta_{0}\right)$, the complete distribution of block sizes is explicitly known, see [24, Theorem 1], from which one can derive some properties of the block of 1 due to exchangeability. For the Bolthausen-Sznitman coalescent ( $\Lambda$ the uniform distribution on $[0,1]$ ) the block of 1 can be characterised as in [32, Corollary 16]. For instance, $f_{1}$ is Markovian for the Bolthausen-Sznitman coalescent.

Different specific aspects of the block of 1 have been analysed for different $\Lambda / \Xi$ -$n$-coalescents including their asymptotics for $n \rightarrow \infty$.

- External branch length: The waiting time for the first jump of the block of 1 in the $n$-coalescent, see e.g. [6-8, 13, 22, 28].
- Minimal clade size: The size $M_{n}$ of the block of 1 for the $n$-coalescent at its first jump. For Kingman's $n$-coalescent and for $\Lambda$ beta-distributed with parameters $(2-\alpha, \alpha)$ with $\alpha \in(1,2), X_{n}$ converges in distribution for $n \rightarrow \infty$, see [6] and [34]. For the Bolthausen-Sznitman $n$-coalescent, $\log \left(M_{n}\right) / \log (n)$ converges in distribution [14]. These results do not cover $\Lambda / \Xi$-coalescents with dust.
- The number of blocks involved in the first merger of the block of 1 , see [34]. The results cover $\Lambda$-coalescents with dust.
- The number of blocks involved in the last merger of the block of 1 , see $[1,2$, 19, 17, 23, 29].
- The small-time behaviour of the block of 1 , see [5, 34].

Due to the exchangeability of the $\Xi$-coalescent, any result for the distribution of the block of 1 holds true for the block containing any other $i \in \mathbb{N}$. We want to further describe $f_{1}$ for $\Xi$-coalescents with dust. For any finite measure $\Xi$ on $\Delta$ which fulfils (2), we introduce

$$
\begin{equation*}
\gamma:=\frac{\Xi(\Delta)}{\mu_{-1}} . \tag{4}
\end{equation*}
$$

We see that $\gamma \in(0,1]$, since

$$
0<\Xi(\Delta)=\int_{\Delta}(x, x) v_{0}(d x) \leq \int_{\Delta}|x| v_{0}(d x)=\mu_{-1}<\infty .
$$

Define $\Delta_{f}:=\bigcup_{k \in \mathbb{N}}\left\{x \in \Delta: x_{1}+\cdots+x_{k}=1\right\}$. We extend Proposition 1 for $\Xi$-coalescents with dust which stay infinite, i.e. have almost surely infinitely many blocks for each $t \geq 0$ (equivalent to $\Xi\left(\Delta_{f}\right)=0$, see Lemma 4). While the extension to $\Xi$-coalescents and the explicit waiting time distributions are a direct follow-up from Pitman's proof, we provide a more detailed description of the jump heights of $f_{1}$. Proposition 1 ensures that the jumps of $f_{1}$ are separated by (almost surely) positive waiting times, we denote the value of $f_{1}$ at its $k$ th jump with $f_{1}[k]$ for $k \in \mathbb{N}$.
Theorem 1. In any $\Xi$-coalescent $\Pi$ with dust and $\Xi\left(\Delta_{f}\right)=0$, the asymptotic frequency process $f_{1}:=\left(f_{1}(t)\right)_{t \geq 0}$ of the block of 1 , defined by Eq. (1), is an increasing pure jump process with càdlàg paths, $f_{1}(0)=0$ and $\lim _{t \rightarrow \infty} f_{1}(t)=1$,
but $f_{1}(t)<1$ for $t>0$ almost surely. The waiting times between almost surely infinitely many jumps are distributed as independent $\operatorname{Exp}\left(\mu_{-1}\right)$ random variables. Its jump chain $\left(f_{1}[k]\right)_{k \in \mathbb{N}}$ can be expressed via stick-breaking

$$
\begin{equation*}
f_{1}[k]=\sum_{i=1}^{k} X_{i} \prod_{j=1}^{i-1}\left(1-X_{j}\right) \tag{5}
\end{equation*}
$$

where $\left(X_{j}\right)_{j \in \mathbb{N}}$ are pairwise uncorrelated, $X_{j}>0$ almost surely and $E\left(X_{j}\right)=\gamma$ for all $j \in \mathbb{N}$. In particular, $E\left(f_{1}[k]\right)=1-(1-\gamma)^{k}$. In general, $\left(X_{j}\right)_{j \in \mathbb{N}}$ are neither independent nor identically distributed.
Remark 1. From Theorem 1, the dependence between $f_{1}$ and its jump times is readily seen as follows. Recall [32, Eq. (51)] that $E\left(f_{1}(t)\right)=1-e^{-t}$ for any $\Lambda$-coalescent with $\Lambda([0,1])=1$. If we would have independence, integrating $E\left(f_{1}(t)\right)$ over the waiting time distribution $\operatorname{Exp}\left(\mu_{-1}\right)$ for the first jump of $f_{1}$ would yield $E\left(f_{1}[1]\right)=$ $\left(1+\mu_{-1}\right)^{-1}$, in contradiction to $E\left(f_{1}[1]\right)=1 / \mu_{-1}$ by Theorem 1 .

Dirac coalescents $\left(\Lambda=\delta_{p}\right.$ for some $\left.p \in(0,1]\right)$ are a family of $\Lambda$-coalescents with dust. They have been introduced as simplified models for populations in species with skewed offspring distributions (reproduction sweepstakes), see [9]. Their jump chains (discrete time Dirac coalescents) can also arise as large population size limits in conditional branching process models [21, Theorem 2.5].

We further characterise $f_{1}$ as follows, including an explicit formula for its distribution at its first jump.
Proposition 2. Let $\Lambda=\delta_{p}, p \in\left[\frac{1}{2}, 1\right)$ and $q:=1-p . f_{1}$ takes values in the set

$$
\begin{equation*}
\mathcal{M}_{p}:=\left\{\sum_{i \in \mathbb{N}} b_{i} p q^{i-1}: b_{i} \in\{0,1\}, 1 \leq \sum_{i \in \mathbb{N}} b_{i}<\infty\right\} . \tag{6}
\end{equation*}
$$

For $x=\sum_{i \in \mathbb{N}} b_{i} p q^{i-1} \in \mathcal{M}_{p}$, we have

$$
\begin{equation*}
P\left(f_{1}[1]=x\right)=p q^{j-1} \prod_{i \in J \backslash\{j\}} P(Y+i \in J) \prod_{i \in[j-1] \backslash J} P(Y+i \notin J)>0, \tag{7}
\end{equation*}
$$

where $Y \stackrel{d}{=} \operatorname{Geo}(p), J:=\left\{i \in \mathbb{N} \mid b_{i}=1\right\}$ and $j:=\max J$. The process $f_{1}$ is not Markovian whereas its jump chain $\left(f_{1}[k]\right)_{k \in \mathbb{N}}$ is Markovian.

## Remarks 2.

- The law of $f_{1}[1]$ is a discrete measure on $[0,1]$ for Dirac coalescents. Surprisingly different properties arise for different values of $p$. For instance, $\mathcal{M}_{2 / 3}=$ $\left\{\sum_{i \in \mathbb{N}} b_{i} 3^{-i}: b_{i} \in\{0,2\}, 1 \leq \sum_{i \in \mathbb{N}} b_{i}<\infty\right\}$ is a subset of the ternary Cantor set which is nowhere dense in $[0,1]$, whereas $\mathcal{M}_{1 / 2}$, the set of all $x \in[0,1]$ with finite 2-adic expansion, is dense in $[0,1]$.
- We omitted $f_{1}[1]$ for the star-shaped coalescent $\left(\Lambda=\delta_{1}\right)$, since it just jumps from 0 to 1 at time $T \stackrel{d}{=} \operatorname{Exp}(1)$.
- Recall that $f_{1}$ is Markovian for the Bolthausen-Sznitman coalescent in contrast to $f_{1}$ for the Dirac coalescents specified above.

Our key motivation was to provide a more detailed description of the jump chain of $f_{1}$, especially properties of the value $f_{1}[1]$ at the first jump which is the asymptotic frequency of the minimal clade. Theorem 1 provides a first-order limit result for all $\Xi$-coalescents with dust.
Corollary 1. Let $\Pi$ be a $\Xi$-coalescent with dust and $\Pi^{(n)}$ its restriction on [n]. Let $M_{n}$ be the minimal clade size, i.e. the size of the block of 1 at its first merger in $\Pi^{(n)}$. Then, $M_{n} / n \rightarrow f_{1}[1]$ almost surely, $f_{1}[1]>0$ almost surely and $E\left(f_{1}[1]\right)=\gamma$.

Compared to the known results listed above for the minimal clade size for dustfree coalescents, the minimal clade size is much larger asymptotically for $n \rightarrow \infty$ $(O(n)$ compared to $o(n))$.

The law of $f_{1}[1]$ in (7) follows from the following more general description of $f_{1}$ [1] for simple $\Lambda$-coalescents. We introduce, for a finite measure $\Lambda$ on [0,1] with $\mu_{-1}=\int_{0}^{1} x^{-1} \Lambda(d x)<\infty$,

$$
\begin{equation*}
\alpha:=\frac{\mu_{-1}}{\mu_{-2}}=\frac{\int_{0}^{1} x^{-1} \Lambda(d x)}{\int_{0}^{1} x^{-2} \Lambda(d x)} . \tag{8}
\end{equation*}
$$

We have $\alpha \in[0,1]$ since $x^{-1} \leq x^{-2}$ on $(0,1]$ (if $\mu_{-1}<\infty$, we have $\Lambda(\{0\})=0$ ). Additionally, $\alpha>0$ if and only if $\mu_{-2}<\infty$, so if $\Lambda$ characterises a simple coalescent (recall that $\mu_{-2} \geq \mu_{-1}>0$ since we exclude $\Lambda=0$ ).
Proposition 3. Let $\Lambda$ fulfil (3). Then,

$$
\begin{equation*}
f_{1}[1]=\sum_{i=1}^{C} B_{i}^{(C)} P_{i} \prod_{j \in[i-1]}\left(1-P_{j}\right)=\sum_{i \in \mathbb{N}} P_{i} \prod_{j \in[i-1]}\left(1-P_{j}\right) \sum_{k \geq i} B_{i}^{(k)} 1_{\{C=k\}}, \tag{9}
\end{equation*}
$$

where $\left(P_{i}\right)_{i \in \mathbb{N}}$ are i.i.d. with $P_{i} \stackrel{d}{=} \mu_{-2}^{-1} x^{-2} \Lambda(d x)$. We have

$$
\left.P\left(C=k \mid\left(P_{i}\right)_{i \in \mathbb{N}}\right)=P_{k} \prod_{j \in[k-1]}\left(1-P_{j}\right), C \text { is Geo( } \alpha\right) \text {-distributed. }
$$

Given $\left(P_{i}\right)_{i \in \mathbb{N}}, C$ and $\left(B_{i}^{(k)}\right)_{k \in \mathbb{N}, i \in[k]}$ are independent and $\left(B_{i}^{(k)}\right)_{k \in \mathbb{N}, i \in[k]}$ is defined as

$$
\begin{align*}
& P\left(\left(B_{1}^{(j)}, \ldots, B_{j}^{(j)}\right)=b \mid\left(P_{i}\right)_{i \in \mathbb{N}}\right) \\
& \quad=\prod_{i \in J \backslash\{j\}} P\left(I(i) \in J \mid\left(P_{i}\right)_{i \in \mathbb{N}}\right) \prod_{i \in[j-1] \backslash J} P\left(I(i) \notin J \mid\left(P_{i}\right)_{i \in \mathbb{N}}\right) \quad \text { almost surely }, \tag{10}
\end{align*}
$$

where $b=\left(b_{1}, \ldots, b_{j}\right) \in\{0,1\}^{j-1} \times\{1\}, J:=\left\{i \in[j] \mid b_{i}=1\right\}$ and, for each $i \in \mathbb{N}$, $I(i):=\min \left\{j \geq i+1: B_{i}^{(j)}=1\right\}$. We have
(i) $P\left(I(i)=i+k \mid\left(P_{i}\right)_{i \in \mathbb{N}}\right)=P_{i+k} \prod_{l=i+1}^{i+k-1}\left(1-P_{l}\right)$ almost surely for $k \in \mathbb{N}$.
(ii) For any $i \in \mathbb{N}$, $I(i)-i$ is Geo( $\alpha$ )-distributed on $\mathbb{N}$. Given $\left(P_{i}\right)_{i \in \mathbb{N}},(I(i))_{i \in \mathbb{N}}$ are independent.

## Remarks 3.

- The distribution of $C$ is known from [16, Proposition 3.1].
- The distribution of $f_{1}[1]$ for Dirac coalescents with $p>\frac{1}{2}$ has a structure somewhat similar to the Cantor distribution, see e.g. [26] and [18]. The Cantor distribution is the law of $\sum_{i \in \mathbb{N}} B_{i} p q^{i-1}$ for $p \in(0,1)$, where $\left(B_{i}\right)_{i \in \mathbb{N}}$ are i.i.d. Bernoulli variables with success probability $\frac{1}{2}$, whereas in our case $\left(B_{i}\right)_{i \in \mathbb{N}}$ are dependent Bernoulli variables with success probabilities $P\left(B_{i}=\right.$ 1) $=P\left(\sum_{k \geq i} B_{i}^{(k)} 1_{\{C=k\}}=1\right)=p q^{i-1}+\sum_{k>i} p^{2} q^{k-1}=p q^{i-1}(1+q)$, see Eq. (9). The Cantor distribution is a shifted infinite Bernoulli convolution. Infinite Bernoulli convolutions are the set of distributions of $\sum_{i \in \mathbb{N}} \omega_{i}(-1)^{B_{i}}$ with $\omega_{i} \in \mathbb{R}$ for $i \in \mathbb{N}$ satisfying $\sum_{i \in \mathbb{N}} \omega_{i}^{2}<\infty$, see [31, Section 2]. They have been an active field of research since the 1930's, e.g. see [10, 35] and the survey [31].

Our main tool for the proofs is Schweinsberg's Poisson construction of the $\Xi$ coalescent. The article is organised as follows. We recall (properties of) the Poisson construction in Section 2. Section 3 characterises staying infinite for $\Xi$-coalescents with dust. These prerequisites are then used to prove the results for $\Xi$-coalescents with dust in Section 4 and for simple $\Lambda$-coalescents in Section 5.

## 2 Poisson construction of a $\Xi$-coalescent and the block of 1

We recall the construction of a $\Xi$-n-coalescent $\Pi$ from [33]. We are only interested in constructing a $\Xi$-coalescent with dust, which implies $\Xi(\{0\})=0$, see Eq. (2).

Let $\mathcal{P}$ be a Poisson point process on $A=[0, \infty) \times \mathbb{N}_{0}^{\infty}$ with intensity measure

$$
\begin{equation*}
v=d t \otimes \int_{\Delta} \otimes_{n \in \mathbb{N}} P^{(x)} v_{0}(d x), \tag{11}
\end{equation*}
$$

where, for $x \in \Delta, P^{(x)}$ is a probability measure on $\mathbb{N}_{0}$ with $P^{(x)}(\{k\})=x_{k}$ and $P^{(x)}(\{0\})=1-|x|$ (Kingman's paintbox) and $v_{0}$ is defined as in Eq. (2). For $n \in \mathbb{N}$, the restriction $\Pi^{(n)}$ of $\Pi$ to $[n]$ can be constructed by starting at $t=0$ with each $i \in[n]$ in its own block. Then, for each subsequent time $(T=) t$ with a Poisson point $\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right)$, merge all present blocks $i$ (at most $n$ ) with identical $k_{i}>0$, where $i$ is the least element of the block (there are only finitely many points of $\mathcal{P}$ that lead to a merger of blocks in $[n]) . \Pi$ is then pathwise defined by its restrictions $\left(\Pi^{(n)}\right)_{n \in \mathbb{N}}$. From now on we will assume without loss of generality that the $\Xi$-coalescent with dust is constructed via the Poisson process $\mathcal{P}$.

The block of 1 can only merge at Poisson points $P=\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right)$ with $K_{1}>0$. We take a closer look at these Poisson points. We introduce exchangeable $(Q)$ indicators following [32, p.1884]: These are exchangeable Bernoulli variables which are conditionally i.i.d. given a random variable $X$ with distribution $Q$ on $[0,1]$ which gives their success probability. Alternatively, we denote these as exchangeable $(X)$ indicators if we can specify $X$.
Lemma 1. For any finite measure $\Xi$ on $\Delta$ fulfilling (2), $\mathcal{P}$ splits into two independent Poisson processes

$$
\mathcal{P}_{1}:=\left\{\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}: K_{1}>0\right\} \quad \text { and } \quad \mathcal{P}_{2}:=\left\{\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}: K_{1}=0\right\} .
$$

$\mathcal{P}_{1}$ has almost surely finitely many points on any set $[0, t] \times \mathbb{N}_{0}^{\infty}$, thus we can order

$$
\mathcal{P}_{1}=\left(\left(T_{j},\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)\right)_{j \in \mathbb{N}},
$$

where $T_{j}<T_{j+1}$ almost surely for $j \in \mathbb{N}$.
$\left(T_{j}\right)_{j \in \mathbb{N}}$ is a homogeneous Poisson process on $[0, \infty)$ with intensity $\mu_{-1}$.
$\left(\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)_{j \in \mathbb{N}}$ is an i.i.d. sequence in $j$ and $\left(1_{\left\{K_{1}^{(1)}=K_{i}^{(1)}\right\}}\right)_{i \geq 2}$ are exchangeable $(Q)$ indicators with

$$
Q:=\frac{1}{\mu_{-1}} \int_{\Delta} \sum_{i \in \mathbb{N}} x_{i} \delta_{x_{i}} v_{0}(d x),
$$

which is a probability measure on $[0,1]$. For $X \stackrel{d}{=} Q$, we have $X>0$ almost surely and $E(X)=\gamma$.

Proof. $\mathcal{P}_{1}$ and $\mathcal{P}_{2}$ are obtained by restricting $\mathcal{P}$ on the disjoint subsets $A_{1}:=[0, \infty) \times$ $\mathbb{N} \times \mathbb{N}_{0}^{\infty}$ and $A_{2}:=[0, \infty) \times\{0\} \times \mathbb{N}_{0}^{\infty}$ of $A$. Thus, $\mathcal{P}_{1}$ and $\mathcal{P}_{2}$ are independent Poisson processes (restriction theorem [25, p.17]) with intensity measures $\nu_{1}=v\left(\cdot \cap A_{1}\right)$ and $\nu_{2}=v\left(\cdot \cap A_{2}\right)$. For any Borel set $B \subseteq[0, \infty)$ and $\lambda$ being the Lebesgue measure,

$$
\begin{equation*}
\nu_{1}\left(B \times \mathbb{N}_{0}^{\infty}\right)=\lambda(B) \int_{\Delta} \underbrace{P^{(x)}(\mathbb{N})}_{=|x|} \prod_{i \geq 2} \underbrace{P^{(x)}\left(\mathbb{N}_{0}\right)}_{=1} \nu_{0}(d x)=\lambda(B) \mu_{-1} \tag{12}
\end{equation*}
$$

Thus, on any bounded set $B, \mathcal{P}_{1}$ has almost surely finitely many points, which can be ordered as described. Projecting $\mathcal{P}_{1}$ on the first coordinate $t$ of $A$ yields a Poisson process with intensity measure $\mu_{-1} d t$ (mapping theorem [25, p.18]).

Now, we project the points of $\mathcal{P}_{1}$ on the coordinate of $\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}$. Recall the construction of a Poisson process as a collection of i.i.d. variables with distribution $(\mu(C))^{-1} \mu$ on sets of finite mass $C$ of the intensity measure $\mu$, e.g. [25, p.23]. It shows that we can treat the collection of $\left(T_{j},\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)$ with, for instance, $T_{i} \in$ $[k, k+1)$ for $k \in \mathbb{N}$ as a random number of i.i.d. random variables with distribution $\left(1 \cdot \mu_{-1}\right)^{-1} \nu_{1}$. Since $\nu_{1}$ has a product structure on $A_{1}$, we have that $\left(\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)_{j \in \mathbb{N}}$ are i.i.d. in $j$ and have distribution, for $m \in \mathbb{N}$,

$$
\begin{equation*}
P\left(\left(K_{i}^{(1)}=l_{i}\right)_{i \in[m]}\right)=\frac{1}{\mu_{-1}} \int_{\Delta} \prod_{i \in[m]} P^{(x)}\left(l_{i}\right) v_{0}(d x)=\frac{1}{\mu_{-1}} \int_{\Delta} \prod_{i \in[m]} x_{l_{i}} v_{0}(d x) \tag{13}
\end{equation*}
$$

for $l_{1} \in \mathbb{N}$ and $l_{2}, \ldots, l_{m} \in \mathbb{N}_{0}$. Consider $\left(1_{\left\{K_{1}^{(1)}=K_{i}^{(1)}\right\}}\right)_{i \geq 2}$. To show that they are exchangeable $(Q)$ indicators, [32, Eq. (27)] has to be fulfilled, i.e. we need to show $P\left(\left\{i \in[m]: K_{i}^{(1)}=K_{1}^{(1)}\right\}=M\right)=E\left(X^{|M|-1}(1-X)^{m-|M|}\right)$ for $X \stackrel{d}{=} Q$ and any $M \subseteq[m]$ with $1 \in M$. Using Eq. (13) we compute

$$
P\left(\left\{i \in[m]: K_{i}^{(1)}=K_{1}^{(1)}\right\}=M\right)=\sum_{j \in \mathbb{N}} P\left(\left\{i \in[m]: K_{i}^{(1)}=j\right\}=M, K_{1}^{(1)}=j\right)
$$

$$
\begin{aligned}
& =\frac{1}{\mu_{-1}} \int_{\Delta} \sum_{j \in \mathbb{N}} x_{j}^{|M|}\left(1-x_{j}\right)^{m-|M|} v_{0}(d x) \\
& =E\left(X^{|M|-1}(1-X)^{m-|M|}\right)
\end{aligned}
$$

Clearly, $P(X>0)=1$ since $\Xi(\{0\})=0$ and $E(X)=\mu_{-1}^{-1} \int_{\Delta}(x, x) \nu_{0}(d x)=$ $\gamma$.

## Remarks 4.

- The properties of the exchangeable( $Q$ ) indicators remind of [32, Lemma 21, Theorem 4] and [33, Proposition 6]. Restricting $\mathcal{P}$ to points with $K_{1}=K_{2}>0$ we can reproduce their results analogously to the proof of Lemma 1.
- $Q$ can be seen as the expected value of the random probability measure $Q_{x}:=$ $|x|^{-1} \sum_{i \in \mathbb{N}} x_{i} \delta_{x_{i}}$ for $x \in \Delta$ with $x$ drawn from $\mu_{-1}^{-1}|x| \nu_{0}(d x)$. In the Poisson construction, this means we draw a "paintbox" $x \in \Delta$ and then record in which box the ball of 1 falls, if we only allow it to fall in boxes $1,2, \ldots$.
- Consider a simple $\Lambda$-coalescent. Projecting $\mathcal{P}_{2}$ on its first component, so $\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \mapsto T$, yields a homogeneous Poisson process with intensity $\mu_{-2}-$ $\mu_{-1}<\infty$. To see this, proceed analogously as for $\mathcal{P}_{1}$. Then, Eq. (12) for $\nu_{2}$ reads the same except for replacing $P^{(x)}(\mathbb{N})$ by $P^{(x)}(\{0\})=1-|x|$.
For a $\Lambda$-coalescent (with $\Lambda(\{0\})=0$ ) the Poisson construction simplifies, since $\Xi$ only has mass on $\left\{x \in \Delta: x_{2}=x_{3}=\cdots=0\right\}$ and thus $\mathcal{P}$ can be seen as a Poisson process on $[0, \infty) \times\{0,1\}^{\infty}$ with intensity measure $d t \otimes \int_{[0,1]} \otimes_{n \in \mathbb{N}} P^{(x)} x^{-2} \Lambda(d x)$, where $P^{(x)}$ is the Bernoulli distribution with success probability $x \in(0,1]$.

When constructing simple $\Lambda$-coalescents, even the process $\mathcal{P}$ itself has almost surely finitely many points $\left(T_{j},\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)$ on any set $[0, t] \times\{0,1\}^{\infty}$ (which we can again order in the first coordinate). As described in [32, Example 19] and analogously to Lemma 1 , we can construct each (potential) merger at point $\left(T_{j},\left(K_{i}^{(j)}\right)_{j \in \mathbb{N}}\right)$ of a simple $\Lambda$-coalescent as follows (while between jumps, we wait independent $\operatorname{Exp}\left(\mu_{-2}\right)$ times $)$. First choose $P_{i} \in(0,1]$ from $\mu_{-2}^{-1} x^{-2} \Lambda(d x)$, we have $E\left(P_{i}\right)=$ $\mu_{-2}^{-1} \int_{[0,1]} x^{-1} \Lambda(d x)=\alpha$. Then, throw independent coins $\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}$ with probability $P_{i}$ for 'heads' (=1) for each block present and merge all blocks whose coins came up 'heads'. Again, $\left(P_{i}\right)_{i \in \mathbb{N}}$ are i.i.d. and the 'coins' $K_{i}^{(j)}$ are exchangeable $\left(P_{i}\right)$ indicators. Analogously to above, we thus have
Lemma 2. Let $\Lambda$ be a finite measure on [0, 1] fulfilling (3). For the Poisson process $\mathcal{P}=\left(T_{j},\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)_{j \in \mathbb{N}},\left(\left(K_{i}^{(j)}\right)_{i \in \mathbb{N}}\right)_{j \in \mathbb{N}}$ is an i.i.d. sequence (in $j$ ) of sequences of exchangeable $\left(P_{j}\right)$ indicators, where $\left(P_{j}\right)_{j \in \mathbb{N}}$ are i.i.d. with $P_{1} \stackrel{d}{=} \mu_{-2}^{-1} x^{-2} \Lambda(d x)$. In particular, $E\left(P_{i}\right)=\alpha$.

Since many proofs will build on the properties of different sets of exchangeable indicators, we collect some well-known properties in the following
Lemma 3. Let $\left(K_{i}\right)_{i \in \mathbb{N}}$ be exchangeable $(X)$ indicators.
a) We have $\lim _{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^{n} K_{i}=X$ almost surely. $X$ is almost surely unique.
b) If $\left(K_{i}\right)_{i \in \mathbb{N}}$ is independent of a $\sigma$-field $\mathcal{F}, X$ is, too.
c) Let $\left(L_{i}\right)_{i \in \mathbb{N}}$ be exchangeable $(Y)$ indicators, independent of $\left(K_{i}\right)_{i \in \mathbb{N}}$. Then, $\left(K_{i} L_{i}\right)_{i \in \mathbb{N}}$ are exchangeable $(X Y)$ indicators and $X, Y$ are independent.

Proof. These properties essentially follow from the de Finetti representation of an infinite series of exchangeable variables as conditionally i.i.d. variables. The lemma is a collection of well-known properties as e.g. described in [3, Sections 2 and 3], arguments of which we use in the following.

An infinite exchangeable sequence is conditionally i.i.d. given an almost surely unique random measure $\alpha$. This measure is the weak limit of the empirical measures, in our case, $n^{-1} \sum_{i=1}^{n} \delta_{K_{i}}$, which has limit $X^{\prime} \delta_{1}+\left(1-X^{\prime}\right) \delta_{0}$ for some random variable $X^{\prime}$ with values in $[0,1]$. Given $\alpha$, the indicators are $\alpha$-distributed. However, since $X$ gives the success probability of each Bernoulli coin, we have $X=X^{\prime}$ almost surely, so $X$ is almost surely unique. The rest of a) is just the strong law of large numbers e.g. from $[3,2.24]\left(E\left(K_{1}\right) \leq 1\right)$, the limit is $X^{\prime}$. Part b) follows from measure theory since the limit is measurable in the $\sigma$-field spanned by the summed variables. For c), we again check Pitman's condition [32, Eq. 27]. Let $M \subseteq[m]$. We have that $X, Y$ are independent from b). With $P\left(K_{i}=L_{i}=1 \mid X, Y\right)=X Y$ almost surely,

$$
\begin{aligned}
P\left(\left\{i \in[m]: K_{i} L_{i}=1\right\}=M\right) & =E\left(P\left(\left\{i \in[m]: K_{i} L_{i}=1\right\}=M \mid X, Y\right)\right) \\
& =E\left((X Y)^{|M|}(1-X Y)^{m-|M|}\right)
\end{aligned}
$$

since given $X, Y$, both $\left(K_{i}\right)_{i \in \mathbb{N}}$ and $\left(L_{i}\right)_{i \in \mathbb{N}}$ are independent. This shows c).

## 3 When does a $\boldsymbol{\Xi}$-coalescent with dust stay infinite?

A crucial assumption for our results is that the $\Xi$-coalescent $\Pi$ has almost surely infinitely many blocks that may merge in the mergers where 1 participates in. The property

$$
P\left(\Pi_{t} \text { has infinitely many blocks } \forall t>0\right)=1
$$

is called staying infinite, while $P\left(\Pi_{t}\right.$ has finitely many blocks $\left.\forall t>0\right)=1$ is the property of coming down from infinity. These properties have been thoroughly discussed for $\Xi$-coalescents, see e.g. [33, 27] and [20].

We recall the condition for $\Xi$-coalescents with dust to stay infinite.
Lemma 4. Let $\Delta_{f}:=\left\{x \in \Delta: x_{1}+\cdots+x_{k}=1\right.$ for some $\left.k \in \mathbb{N}\right\}$ and $\Xi$ be a finite measure on $\Delta$ fulfilling Eq. (2). The $\Xi$-coalescent stays infinite if and only if $\Xi\left(\Delta_{f}\right)=0$. If $\Xi\left(\Delta_{f}\right)>0$, then the $\Xi$-coalescent has infinitely many blocks until the first jump of $f_{1}$ almost surely and the $\Xi$-coalescent neither comes down from infinity nor stays infinite.

Proof. Let $\Delta^{*}:=\{x \in \Delta:|x|=1\}$. All $\Xi$-coalescents considered are constructed via the Poisson construction with Poisson point process $\mathcal{P}$.

First, assume $\Xi\left(\Delta^{*}\right)=0$. We recall the (well-known) property that for a $\Xi$ coalescent with dust $\Xi\left(\Delta^{*}\right)=0$ is equivalent to $P\left(S_{t}>0 \forall t\right)=1$, where $S_{t}$ is the asymptotic frequency of the dust component. We use the remark on [12, p.1091]: For $\Xi$-coalescents with dust, $\left(-\log S_{t}\right)_{t \geq 0}$ is a subordinator. The subordinator jumps to $\infty$ (corresponds to $S_{t}=0$ ) if and only if for its Laplace exponent $\Phi$, we have
$\lim _{\eta \backslash 0} \Phi(\eta)>0$. For a $\Xi$-coalescent with dust we have $\lim _{\eta \backslash 0} \Phi(\eta)=\int_{\Delta^{*}} \nu_{0}(d x)$. Hence, $\Xi\left(\Delta^{*}\right)=0$ almost surely guarantees infinitely many singleton blocks for all $t \geq 0$, so the corresponding $\Xi$ coalescent stays infinite.

Now assume $\Xi\left(\Delta^{*}\right)>0$. The subordinator $\left(-\log S_{t}\right)_{t \geq 0}$ jumps from finite values $\left(S_{t}>0\right)$ to $\infty\left(S_{t}=0\right)$ after an exponential time with rate $\nu_{0}\left(\Delta^{*}\right)$. This shows that the $\Xi$-coalescent does not come down from infinity. Assume further that $\Xi\left(\Delta_{f}\right)=0$. Then, [33, Lemma 31] shows that the $\Xi$-coalescent either comes down from infinity or stays infinite, so it stays infinite.

Finally, assume $\Xi\left(\Delta_{f}\right)>0$. Split $\mathcal{P}$ into independent Poisson processes $\mathcal{P}_{1}^{\prime}:=$ $\left\{\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}: \kappa \in \Delta_{f}\right\}$ and $\mathcal{P}_{2}^{\prime}:=\left\{\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}: \kappa \notin \Delta_{f}\right\}$, where $\kappa:=\left(\lim _{n \rightarrow \infty} n^{-1} \sum_{i \in[n]} 1_{\left\{K_{i}=j\right\}}\right)_{j \in \mathbb{N}}$ (again restriction theorem [25, p.17], Lemma 3 shows $\kappa$ exists almost surely). Their intensity measures are defined as in Eq. (11), but using $\nu_{1}^{\prime}(\cdot):=\nu_{0}\left(\cdot \cap \Delta_{f}\right)$ and $\nu_{2}^{\prime}:=\nu_{0}-v_{1}^{\prime}$ instead of $\nu_{0}$. Since $\nu_{1}^{\prime}\left(\Delta_{f}\right) \leq$ $\mu_{-1}<\infty$, for any $t>0$ there are almost surely finitely many $P \in \mathcal{P}_{1}^{\prime}$ with $T<t$. Consider such $P=\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right)$ with $T$ smallest. Observe that until $T$, we can construct the $\Xi$-coalescent using only the points of $\mathcal{P}_{2}^{\prime}$, which is the construction of a $\Xi^{\prime}$-coalescent with $\Xi^{\prime}(d x):=(x, x) v_{2}^{\prime}(d x)$. Since $\int_{\Delta}|x| v_{2}^{\prime}(d x)<\mu_{-1}<\infty$ and $\Xi^{\prime}\left(\Delta_{f}\right)=0$, the proof steps above show that the $\Xi$-coalescent has infinitely many blocks until $T$. Now consider the merger at time $T$. The form of $v_{1}^{\prime}$ ensures that $\left(K_{i}\right)_{i \in \mathbb{N}}$ can only take finitely many values, and Lemma 3a) ensures that infinitely many $K_{i}$ 's show each value. Thus, all blocks present before time $T$ are merged at $T$ into a finite number of blocks (given by which $K_{i}$ 's show the same number). This shows that if $\Xi\left(\Delta_{f}\right)>0$, the $\Xi$-coalescent stays neither infinite nor comes down from infinity. Additionally, this shows that either the block of 1 already merged at least once before $T$ or it merges at $T$, thus there are infinitely many blocks before the first merger of 1.

## 4 The block of 1 in $\Xi$-coalescents with dust - proofs and remarks

Proof of Theorem 1. As in Lemma 1, split the Poisson point process $\mathcal{P}$ used to construct the $\Xi$-coalescent in $\mathcal{P}_{1}$ and $\mathcal{P}_{2}$. We also use the notation from Lemma 1 and its proof. The block of 1 in the $\Xi$ - $n$-coalescent for any $n \in \mathbb{N}$ can only merge at times $t$ for which there exists a Poisson point $\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}_{1}$. Lemma 1 states that the set of times $T$ forms a homogeneous Poisson process with rate $\mu_{-1}$. This shows that potential jump times are separated by countably many independent $\operatorname{Exp}\left(\mu_{-1}\right)$ random variables. Kingman's correspondence yields that $f_{1}$ exists almost surely at each potential jump time. To see this, observe that even though the partition of $\mathbb{N}$ induced by the Poisson construction is not exchangeable, the partition on $\mathbb{N} \backslash\{1\}$ is, and the asymptotic frequencies of the former and the latter coincide. Since $f_{1}$ is by definition constant between these jump points, $f_{1}$ has càdlàg paths almost surely. Since any blocks change by mergers, $f_{1}$ is increasing.

The value of $f_{1}$ at 0 follows by definition. Since $\Pi$ stays infinite (see Lemma 4), at each $P \in \mathcal{P}_{1}$ infinitely many blocks can potentially merge. Lemma 1 shows that the indicators of whether blocks present immediately before $P$ merge with the block of 1 are exchangeable $(X)$ indicators with $X>0$ almost surely. Then, Lemma 3 ensures that a positive fraction of them almost surely does, causing $f_{1}$ to jump (since
a positive fraction of merging blocks has positive frequency). Thus, every Poisson point leads to a merger almost surely, which shows that $f_{1}$ jumps at all potential jump times described above. Since, for all $t$, either $S_{t}>0$ or non-dust blocks not including 1 exist (having asymptotical frequency $>0$ ), $f_{1}(t)<1$ for all $t \geq 0$.

We consider the jump chain of $f_{1}$. Set $X_{1}:=f_{1}[1]$. Since $f_{1}[k] \in(0,1)$ for all $k \in \mathbb{N}$ and $f_{1}$ increases, $f_{1}[k+1]=f_{1}[k]+\left(1-f_{1}[k]\right) X_{k+1}$ for $X_{k+1} \in(0,1)$. Iterating this yields $f_{1}[k]=\sum_{i=1}^{k} X_{i} \prod_{j=1}^{i-1}\left(1-X_{j}\right)$ for $k \geq 2$. The properties of $\left(X_{k}\right)_{k \in \mathbb{N}}$ follow from the Poisson construction and Lemma 1. Consider the blocks present at time $T_{k}-$, where the $k$ th Poisson point of $\mathcal{P}_{1}$ is $P_{k}=\left(T_{k},\left(K_{i}^{(k)}\right)_{i \in \mathbb{N}}\right)$. The block with least element $i$ merges with the block of 1 if $K_{i}^{(k)}=K_{1}^{(k)}$. Consider the $k$ th Poisson point at time $T_{k} . X_{k}$ gives the fraction of the asymptotic frequency of nonsingleton blocks and singleton blocks at time $T_{k}-$, i.e. the fraction of $1-f_{1}\left(T_{k-}\right)$, that is merged with the block of 1 at $T_{k}$. Denote $L_{i}^{(k-)}:=1_{\{\{i\}}$ is a block at $\left.T_{k}-\right\}$. Then, recording the asymptotic frequencies of merged non-singleton and singleton blocks,

$$
X_{k}=\frac{1}{1-f_{1}\left(T_{k}-\right)}\left(\sum_{i \geq 2} 1_{\left\{K_{1}^{(k)}=K_{i}^{(k)}\right\}} f_{i}\left(T_{k}-\right)+\lim _{n \rightarrow \infty} \frac{1}{n} \sum_{i=2}^{n} 1_{\left\{K_{1}^{(k)}=K_{i}^{(k)}\right\}} L_{i}^{(k-)}\right)
$$

Since by construction, $\Pi_{T_{k}-} \backslash\{1\}$ is an exchangeable partition of $\mathbb{N} \backslash\{1\},\left(L_{i}^{(k-)}\right)_{i \in \mathbb{N}}$ are exchangeable $\left(S_{t-}\right)$ indicators with $S_{t-}=1-\sum_{i=1}^{\infty} f_{i}\left(T_{k}-\right)$. Recall that Lemma 1 tells us that $\left(1_{\left\{K_{1}^{(k)}=K_{i}^{(k)}\right\}}\right)_{i \geq 2}$ are exchangeable $\left(X^{\prime}\right)$ indicators with $X^{\prime} \stackrel{d}{=} Q .\left(K_{i}^{(k)}\right)_{i \in \mathbb{N}}$ is independent from $\left(\Pi_{t}\right)_{t<T_{k}}$, since the Poisson points of $\mathcal{P}_{1}$ are i.i.d., so Lemma 3 c) and a) show

$$
\begin{equation*}
X_{k} \stackrel{\text { a.s. }}{=} \sum_{i \geq 2} 1_{\left\{K_{1}^{(k)}=K_{i}^{(k)}\right\}} \frac{f_{i}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)}+X^{\prime} \frac{1-\sum_{i=1}^{\infty} f_{i}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)} . \tag{14}
\end{equation*}
$$

The independence of $\left(K_{i}^{(k)}\right)_{i \in \mathbb{N}}$ from $\left(\Pi_{t}\right)_{t<T_{k}}$ is also crucial for the next two equations. Compute, with $P\left(K_{1}^{(k)}=K_{i}^{(k)}\right)=E\left(X^{\prime}\right)=\gamma$ for $i \in \mathbb{N}$,

$$
\begin{aligned}
E\left(X_{k}\right) & =\sum_{i \geq 2} P\left(K_{1}^{(k)}=K_{i}^{(k)}\right) E\left(\frac{f_{i}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)}\right)+E\left(X^{\prime}\right) E\left(\frac{1-\sum_{i=1}^{\infty} f_{i}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)}\right) \\
& =\gamma E\left(\frac{1-f_{1}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)}\right)=\gamma
\end{aligned}
$$

Analogously, for $l<k, X_{l}$ only depends on Poisson points $P_{1}, \ldots, P_{l}$, so

$$
\begin{aligned}
E\left(X_{k} X_{l}\right) & =E\left(\sum_{i \geq 2} 1_{\left\{K_{1}^{(k)}=K_{i}^{(k)}\right\}} \frac{f_{i}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)} X_{l}+X^{\prime} \frac{1-\sum_{i=1}^{\infty} f_{i}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)} X_{l}\right) \\
& =\gamma E\left(\frac{1-f_{1}\left(T_{k}-\right)}{1-f_{1}\left(T_{k}-\right)} X_{l}\right)=\gamma^{2}
\end{aligned}
$$

showing that $X_{k}, X_{l}$ are uncorrelated. An analogous computation shows that $E\left(\prod_{i \in\left\{l_{1}, \ldots, l_{m}\right\}} X_{l_{i}}\right)=\prod_{i \in\left\{l_{1}, \ldots, l_{m}\right\}} E\left(X_{l_{i}}\right)$ for distinct $l_{1}, \ldots, l_{m} \in \mathbb{N}$. With this,

$$
E\left(f_{1}[k]\right)=\sum_{i=1}^{k} E\left(X_{i}\right) \prod_{j=1}^{i-1}\left(1-E\left(X_{j}\right)\right)=\sum_{i=1}^{k} \gamma(1-\gamma)^{i-1}=1-(1-\gamma)^{k} .
$$

To prove $\lim _{t \rightarrow \infty} f_{1}(t)=1$ almost surely, observe that $f_{1}$ is bounded and increasing, thus $\lim _{t \rightarrow \infty} f_{1}(t)$ exists. Monotone convergence and $\lim _{t \rightarrow \infty} E\left(f_{1}(t)\right)=$ $\lim _{k \rightarrow \infty} E\left(f_{1}[k]\right)=1$ show the desired. Note that $\left(X_{k}\right)_{k \in \mathbb{N}}$ is in general neither independent nor identically distributed, see Section 6.

Proof of Corollary 1. By the Poisson construction the block of 1 for $\Pi^{(n)}$ can only merge at times given by Poisson points in $\mathcal{P}_{1}$. Consider $\left(T_{1},\left(K_{i}^{(1)}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}_{1}$. While $T_{1}$ is the time of the first jump of $f_{1}$ (see the proof of Theorem 1 ), there does not necessarily need to be a merger of $\{1\}$ in the $n$-coalescent $\Pi^{(n)}$, if we have $K_{1}^{(1)} \neq K_{i}^{(1)}$ for the least elements $i$ of all other blocks of $\Pi^{(n)}$ immediately before $T_{1}$. However, Lemma 1 shows that $\left(1_{\left\{K_{1}^{(1)}=K_{i}^{(1)}\right\}}\right)_{i \geq 2}$ are exchangeable indicators. The mean $n^{-1} \sum_{i=2}^{n} 1_{\left\{K_{1}^{(1)}=K_{i}^{(1)}\right\}}$, as argued in the proof of Theorem 1, converges to an almost surely positive random variable for $n \rightarrow \infty$. As shown in Lemma 4 , any $\Xi$-coalescent with dust has infinitely many blocks almost surely before $T_{1}$. Thus, there exists $N$, a random variable on $\mathbb{N}$, so that 1 is also merging at time $T_{1}$ in $\Pi^{(n)}$ for $n \geq N$ almost surely. This yields $\lim _{n \rightarrow \infty} n^{-1} M_{n}=\lim _{n \rightarrow \infty} n^{-1}\left|B_{1}\left(T_{1}\right) \cap[n]\right|=f_{1}\left(T_{1}\right)=f_{1}[1]$ almost surely. All further claims follow from Theorem 1.

Remark 5. Let $Q^{(n)}$ be the number of blocks merged at the first collision of the block of 1 in a $\Lambda$ - $n$-coalescent with dust. [34, 1.4] shows that $n^{-1} Q^{(n)}$ converges in distribution. We argue that this convergence also holds in $L^{p}$ for all $p>0$ and, for simple $\Lambda$-n-coalescents, almost surely.

The proof of Corollary 1 shows that $\left(T_{1},\left(K_{i}^{(1)}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}_{1}$ causes the first merger in the $n$-coalescent for $n$ large enough (almost surely, but since $n^{-1} Q^{(n)} \in[0,1]$ for all $n$, convergence in $L^{p}$ is not affected by the null set excluded). Split $Q^{(n)}$ into $Q_{0}^{(n)}$, the number of non-singleton blocks and $Q_{1}^{(n)}$, the number of singleton blocks merged at $T_{1}$. For the limit, we can ignore the non-singleton blocks merged. To see this, recall $Q_{0}^{(n)} \leq K_{n}$, where $K_{n}$ is the total number of mergers for the $\Lambda$-n-coalescent, since a non-singleton block has to be the result of a merger. [12, Lemma 4.1] tells us that $n^{-1} K_{n} \rightarrow 0$ in $L^{1}$ for $n \rightarrow \infty$ for $\Xi$-coalescents with dust. This shows that the $L^{1}$ limit of $n^{-1} Q^{(n)}$ is the same as of the one of $n^{-1} Q_{1}^{(n)} \cdot n^{-1} Q_{1}^{(n)}$ already appeared in the part of the proof of Theorem 1 leading to Eq. (14), its limit almost surely exists and equals $X^{\prime} \frac{1-\sum_{i=1}^{\infty} f_{i}\left(T_{1}-\right)}{1-f_{1}\left(T_{1}-\right)}$. Since $n^{-1} Q_{1}^{(n)}$ is bounded in $[0,1]$, it also converges in $L^{p}$, $p>0$. So $n^{-1} Q^{(n)}$ converges in $L^{1}$. Since it is bounded in $[0,1]$ it also converges in $L^{p}, p>0$. For simple $\Xi$-n-coalescents, [11, Lemma 4.2] shows $n^{-1} K_{n} \rightarrow 0$ almost surely, so in this case the steps above ensure also almost sure convergence of $n^{-1} Q^{(n)}$.

## 5 The block of 1 in simple $\boldsymbol{\Lambda}$-coalescents - proofs and remarks

Proof of Proposition 3. Let $\mathcal{P}:=\left(P_{i}\right)_{i \in \mathbb{N}}$ be the coin probabilities coming from the Poisson process used to construct the simple $\Lambda$-coalescent $\Pi$ as described in Section 2. As shown in the proof of Theorem 1, the Poisson point belonging to $P_{C}$ where 1 first throws 'heads' in the Poisson construction is the Poisson point where $f_{1}$ jumps for the first time. We have $P(C=k \mid \mathcal{P})=P_{k} \prod_{i=1}^{k-1}\left(1-P_{i}\right)$. Integrating the condition and using the independence of $\left(P_{i}\right)_{i \in \mathbb{N}}$ as well as $E\left(P_{1}\right)=\alpha$ (see Lemma 2), we see that $C$ is geometrically distributed with parameter $\alpha$.

To describe $f_{1}[1]$ at the $C$ th merger (Poisson point), recall that the restriction $\Pi_{-1}$ of $\Pi$ to $\mathbb{N} \backslash\{1\}$ has the same asymptotic frequencies as $\Pi$. Thus, we can see $f_{1}[1]$ as the asymptotic frequency of the newly formed block of $\Pi_{-1}$ at the time of the Poisson point $P_{C}$. This follows since $\Pi_{-1}$ has infinitely many blocks before (see Lemma 4) and then, as in the proof of Theorem 1, there will be a newly formed block of $\Pi_{-1}$ at the $C$ th Poisson point (and the unrestricted block in $\Pi$ includes 1).

We consider $\Pi_{-1}$ at the $k$ th Poisson point with coin probability $P_{k}$. For $\{i\} \in$ $\mathbb{N} \backslash\{1\}$ to remain a (singleton) block and not be merged for the first $k-1$ mergers and then to be merged at the $k$ th, we need $\prod_{j \in[k-1]}\left(1-K_{i}^{(j)}\right)=1$ and $K_{i}^{(k)}=1$. $\left(1_{\left\{\prod_{j \in[k-1]}^{\left.\left(1-K_{i}^{(j)}\right)=1, K_{i}^{(k)}=1\right\}}\right.}\right)_{i \in \mathbb{N}}$ are exchangeable $\left(P_{k} \prod_{j \in[k-1]}\left(1-P_{j}\right)\right)$ indicators. Let

$$
\mathcal{S}_{k}=\left\{i \in \mathbb{N} \backslash\{1\}: \prod_{j \in[k-1]}\left(1-K_{i}^{(j)}\right)=1, K_{i}^{(k)}=1\right\}
$$

be the set of $i \in \mathbb{N} \backslash\{1\}$ whose first merger is the $k$ th overall merger. We call $\mathcal{S}_{k}$ the $k$ th singleton set (of $\Pi_{-1}$ ). From the strong law of large numbers for exchangeable indicators, see Lemma 3a), we directly have that $\mathcal{S}_{k}$ has asymptotic frequency $P_{k} \prod_{j \in[k-1]}\left(1-P_{j}\right)$ almost surely.

Now, consider the asymptotic frequency $f^{*}[k]$ of the newly formed block at the $k$ th merger of $\Pi_{-1}$. By construction, there is only one newly formed block at each merger. $\mathcal{S}_{k}$ is a part of the newly formed block. Any other present block with more than two elements (non-singleton block) is merged if and only if its indicator $K_{i}^{(k)}=$ 1 (we order by least elements). For $k=1$, the newly formed block is $\mathcal{S}_{1}$. For $k=2$, it is either $\mathcal{S}_{2}$ or $\mathcal{S}_{1} \cup \mathcal{S}_{2}$, if the coin of the the block $\mathcal{S}_{1}$ formed in the first merger comes up 'heads'.

Applied successively, this shows that the newly formed block at the $k$ th merger consists of a union of a subset of the singleton sets $\left(\mathcal{S}_{k^{\prime}}\right)_{k^{\prime}<k}$ and the set $\mathcal{S}_{k}$. For its asymptotic frequency, we have

$$
\begin{equation*}
f^{*}[k]=\sum_{i=1}^{k} B_{i}^{(k)} P_{i} \prod_{j \in[i-1]}\left(1-P_{j}\right)>0 \tag{15}
\end{equation*}
$$

where the $B_{i}^{(k)}, i \in[k]$, are non-independent Bernoulli variables which are 1 if the $i$ th singleton set $\mathcal{S}_{i}$ is a part of the newly formed block at the $k$ th merger of $\Pi_{-1}$.

If $\Lambda(\{1\})>0, P_{k}=1$ is possible. In this case, at the $k$ th Poisson point all remaining singletons form $\mathcal{S}_{k}$ and all blocks present at merger $k-1$ merge with $\mathcal{S}_{k}$.

There are no mergers at Poisson points $P_{l}, l>k$, so we do not consider Eq. (15) for $l>k$.

We have $f_{1}[1]=f^{*}[C]$. Given $\mathcal{P},\left(f^{*}[k]\right)_{k \in \mathbb{N}}$ is independent of $C$. Thus, Eq. (9) is implied by Eq. (15).

Assume $\Lambda(\{1\})=0$. For $\left(B_{i}^{(k)}\right)_{k \in \mathbb{N}, i \in[k]}$, we have $B_{k}^{(k)}=1$ for all $k \in \mathbb{N}$ since the $k$ th singleton set is formed at the $k$ th Poisson point and is a part of the newly formed block. The coins thrown at the $k$ th Poisson point to decide whether other singleton sets $\mathcal{S}_{i}, \mathcal{S}_{j}$ with $i, j<k$ are also parts of the newly formed block are either independent given $\mathcal{P}$ when they are in different blocks, or identical when they are in the same block. The set $\mathcal{S}_{i}$ uses the coin of the block newly formed at the $i$ th merger. Let $I(i)$ be the Poisson point at which this block merges again (and $\mathcal{S}_{i}$ with it). At the $I(i)$ th Poisson point and for all further Poisson points indexed with $j \geq I(i)$, we have $B_{i}^{(j)}=B_{I(i)}^{(j)}$, since the singleton sets $S_{i}$ and $S_{I(i)}$ are in the same block for mergers $j \geq I(i)$.

The property (i) of $I(i)$ in the proposition follow directly from its definition as the minimum number of coin tosses until the first comes up 'heads'. The property (ii) is just integrating (i) and using that $\left(P_{i}\right)_{i \in \mathbb{N}}$ are i.i.d. with $E\left(P_{1}\right)=\alpha$ (see Lemma 2), the conditional independence is the conditional independence of coin tosses of distinct blocks from the Poisson construction. To see Eq. (10), observe that $\mathcal{S}_{i}$ for $i<j$ is a part of the newly formed block at the $j$ th merger of the $\Lambda$-coalescent $(i \in J)$ if and only if $I(i) \in J$. If $I(i) \in J$, either we have $I(i)=j$, so $\mathcal{S}_{i}$ is merged for the first time after it has been formed at the $j$ th merger, or we have that $I(i)<j$ which means that it has already merged with at least one other singleton set and that, as parts of the same block, they both again merged at the $j$ th merger. If $I(i) \notin J$, the singleton set $\mathcal{S}_{i}$ neither merges at the $j$ th merger for the first time after being formed nor merges with any other singleton set before that is then merging at the $j$ th merger, so $\mathcal{S}_{i}$ is not a part of the newly merged block at the $j$ th merger.

If $\Lambda(\{1\})=0$, the arguments hold true for all $j \in \mathbb{N}$. If $\Lambda(\{1\})>0$ this holds true for all $i, j \leq K:=\min _{k \in \mathbb{N}}\left\{P_{k}=1\right\}(<\infty$ almost surely), where all singleton sets merge and $f^{*}[K]=1$. However, in this case $C \leq K$, so we still can establish Eq. (9).

## Remarks 6.

- $(I(i))_{i \in \mathbb{N}}$ is useful to construct the asymptotic frequencies of the $\Lambda$-coalescent. Given $\mathcal{P}$, at the kth merger, there are the singleton sets $\left(\mathcal{S}_{j}\right)_{j \in[k]}$ with almost sure frequencies $P_{j} \prod_{i \in[j-1]}\left(1-P_{i}\right)$ which were already formed in the $k$ collisions, and unmerged singleton blocks with frequency $\prod_{i \in[k]}\left(1-P_{i}\right)$. Using $(I(i))_{i \in[k]}$, we can indicate which singleton sets form a block. $\mathcal{S}_{i}$ is a single block if $I(i)>k$, if $I(i) \leq k$ it is a part of a block where $\mathcal{S}_{I(i)}$ is also a part of. This can be seen as a discrete version of the construction of the $\Lambda$-coalescent from the process of singletons as described in [15, Section 6.1]
- The variables $(I(i))_{i \in \mathbb{N}}$ are useful to express other quantities of the $\Lambda$-coalescent. For instance, the number of non-singleton blocks in a simple $\Lambda$-coalescent at the $k$ th merger is given by $k-\sum_{i \in[k-1]} 1_{\{I(i) \leq k\}}$.
To prove Proposition 2, we need the following result.

Lemma 5. For $p \in\left[\frac{1}{2}, 1\right)$, each $x \in \mathcal{M}_{p}$ from Eq. (6) has a unique representation in $\mathcal{M}_{p}$.

Proof. We adjust the proof of [4, Theorem 7.11]. Assume that $x \in \mathcal{M}_{p}$ has two representations $x=\sum_{i \in \mathbb{N}} b_{i} p q^{i-1}=\sum_{i \in \mathbb{N}} b_{i}^{\prime} p q^{i-1}$ with $b_{i} \neq b_{i}^{\prime}$ for at least one $i$. Let $i_{0}$ be the smallest integer with $b_{i_{0}} \neq b_{i_{0}}^{\prime}$. Without restriction, assume $b_{i_{0}}-b_{i_{0}}^{\prime}=1$. Then,

$$
0=\sum_{i \in \mathbb{N}} b_{i} p q^{i-1}-\sum_{i \in \mathbb{N}} b_{i}^{\prime} p q^{i-1}=p q^{i_{0}-1}+\sum_{i>i_{0}}\left(b_{i}-b_{i}^{\prime}\right) p q^{i-1}
$$

Thus, $p q^{i_{0}-1}=\sum_{i>i_{0}}\left(b_{i}^{\prime}-b_{i}\right) p q^{i-1}<\sum_{i>i_{0}} p q^{i-1}=p q^{i_{0}}$, simplifying to $p<q$, in contradiction to the assumption $p \geq \frac{1}{2}$.
Proof of Proposition 2. From Eq. (15) we see that $f_{1}$ only takes values in $\mathcal{M}_{p}$, since $P_{k}=p$ for all $k \in \mathbb{N}$ and $C<\infty$ almost surely. Recall the definition of the singleton sets $\mathcal{S}_{i}$ and their properties from the proof of Proposition 3. The asymptotic frequency of $\mathcal{S}_{i}$ is $p q^{i-1}$ almost surely. Lemma 5 ensures that there is a unique representation $f_{1}[l]=x=\sum_{i=1}^{\infty} b_{i} p q^{i-1}$ in $\mathcal{M}_{p}$, let $J:=\left\{i \in \mathbb{N}: b_{i}=1\right\}$ and $j:=\max J$. This means that $f_{1}[l]=x$ is equivalent to that the block of 1 at its $l$ th jump consists of the union of all $\mathcal{S}_{i}$ with $i \in J$ and 1 . This also shows that the $l$ th jump of $f_{1}$ is at the $j$ th jump of the Dirac coalescent, since if $f_{1}$ jumps at the $k$ th merger of the Dirac coalescent, the newly formed block includes $\mathcal{S}_{k}$.

Since $P_{i}=p$ for all $i \in \mathbb{N}$, we have $\alpha=p$ and Eq. (9) simplifies to $f_{1}[1]=$ $\sum_{i=1}^{C} B_{i}^{(C)} p q^{i-1}$, where $C \stackrel{d}{=} \operatorname{Geo}(p)$ is independent from $\left(B_{i}^{(k)}\right)_{k \in \mathbb{N}, i \in[k]}$. The latter fulfil

$$
\begin{equation*}
P\left(B_{i}^{(j)}=b_{i} \forall i \in[j-1]\right)=\prod_{i \in J \backslash\{j\}} P(Y+i \in J) \prod_{i \in[j \backslash \backslash J} P(Y+i \notin J) \tag{16}
\end{equation*}
$$

with $Y \stackrel{d}{=} G e o(p)$, since the joint distribution in Proposition 3 again simplifies, we can ignore the conditioning and $I(i)-i \stackrel{d}{=} G e o(p)$ for all $i \in \mathbb{N}$.

Since $f_{1}[1]=x$ uniquely determines the values of $C$ and $\left(B_{i}^{(C)}\right)_{i \in[C]}$, we have

$$
\begin{equation*}
P\left(f_{1}[1]=x\right)=P(C=j) P\left(B_{1}^{(j)}=b_{1}, \ldots, B_{j-1}^{(j)}=b_{j-1}\right) \tag{17}
\end{equation*}
$$

which shows Eq. (7) when we insert the distributions expressed in terms of their geometric distributions.

In order to verify that the jump chain $\left(f_{1}[i]\right)_{i \in \mathbb{N}}$ is Markovian, we show that $f_{1}[1], \ldots, f_{1}[l]$ does not contain more information on $f_{1}[l+1]$ than $f_{1}[l]$ does. Without restriction, assume that the $l$ th jump $f_{1}[l]$ of $f_{1}$ takes place at the $k$ th jump of the Dirac coalescent. Then, $f_{1}[l+1]$ is constructed from the blocks present after the $k$ th merger. For each subsequent Poisson point $P_{k+1}, \ldots$, blocks present are merged if their respective coins come up 'heads' until (and including), at $P_{k^{\prime}}$, the coin of the block of 1 comes up 'heads' for the first time since $P_{k}$. Thus, only information about the block partition at merger $k$ can change the law of the next jump. $f_{1}[l]=x$ gives the information which singleton sets $\mathcal{S}_{1}, \ldots, \mathcal{S}_{k}$ are parts of the block of 1 at merger
$k$ of $\Pi$ and which are not. $f_{1}[l]=x$ contains no information about how the other singleton sets, $\mathcal{S}_{i}$ with $b_{i}=0$, are merged into blocks at collisions before $k$ apart from that it tells us that $B_{i}^{(j)}=0$ for $j \in J$ and $i \notin J$, which means that all $\mathcal{S}_{i}$ with $i \notin J$ did not merge at the $j$ th collisions, $j \in J$. This is due to that any $\mathcal{S}_{i}$ with $B_{i}^{(j)}=1$ would merge with the newly formed block at merger $j$ and thus would be in a block with $\mathcal{S}_{j}$ and also in the block of 1 at merger $k$. However, analogously we see that knowing $f_{1}[1], \ldots, f_{1}[l]$ does not give any additional information about the block structure at the $k$ th merger, but only how the set of $\mathcal{S}_{i}$ which are in the block of 1 at merger $k$ behaved at the earlier mergers $J$. Thus, $\left(f_{1}[l]\right)_{l \in \mathbb{N}}$ is Markovian. However, $\left(f_{1}(t)\right)_{t \geq 0}$ is not Markovian. In order to see this consider, for $0<t_{0}<t_{1}<t_{2}$,

$$
\begin{aligned}
p\left(t_{2}, t_{1}, t_{0}\right): & =P\left(f_{1}\left(t_{2}\right)=p+p q^{2} \mid f_{1}\left(t_{1}\right)=p, f_{1}\left(t_{0}\right)=0\right) \\
& =\frac{P\left(f_{1}\left(t_{2}\right)=p+p q^{2}, f_{1}\left(t_{1}\right)=p, f_{1}\left(t_{0}\right)=0\right)}{P\left(f_{1}\left(t_{1}\right)=p, f_{1}\left(t_{0}\right)=0\right)}
\end{aligned}
$$

We will show that $p\left(t_{2}, t_{1}, t_{0}\right)$ depends on $t_{0}$, which shows that $f_{1}$ is not Markovian.
We can express all events in terms of the independent waiting times for Poisson points, i.e. the successive differences between the first component $T$ of the Poisson points $\left(T,\left(K_{i}\right)_{i \in \mathbb{N}}\right) \in \mathcal{P}$. Here, we use the split of the Poisson points into the independent Poisson point processes $\mathcal{P}_{1}$ and $\mathcal{P}_{2}$ from Lemma 1. The waiting times between points in $\mathcal{P}_{1}$ are $\operatorname{Exp}\left(\mu_{-1}\right)$-distributed, the waiting times between points in $\mathcal{P}_{2}$ are $\operatorname{Exp}\left(\mu_{-2}-\mu_{-1}\right)$-distributed, see Lemma 1 and Remark 4. We will relabel $\tau=\mu_{-1}$ and $\rho=\mu_{-2}-\mu_{-1}$ for a clearer type face. Let $T_{1}, T_{2}, \ldots$ be the waiting times between points in $\mathcal{P}_{1}$ and $T_{1}^{\prime}, T_{2}^{\prime}, \ldots$ be the waiting times between points in $\mathcal{P}_{2}$. All waiting times are independent one from another. We recall that for $T \stackrel{d}{=} \operatorname{Exp}(\alpha)$, $P(T>a)=e^{-\alpha a}$ and $P(T \in(a, a+b])=e^{-\alpha a}\left(1-e^{-\alpha b}\right)$ for $a, b \geq 0$.

The event $\left\{f_{1}\left(t_{1}\right)=p, f_{1}\left(t_{0}\right)=0\right\}$ means that the first jump of $f_{1}$ adds the singleton set $\mathcal{S}_{1}$ at a time in $\left(t_{0}, t_{1}\right]$. Thus, there has to be only a single point of $\mathcal{P}_{1}$ with first component $T_{1} \leq t_{1}$ and the smallest time $T_{1}^{\prime}$ of points of $\mathcal{P}_{2}$ has to be greater than $T_{1}$. We compute, conditioning on $T_{1}$ for the third equation,

$$
\begin{aligned}
P\left(f_{1}\left(t_{1}\right)=p, f_{1}\left(t_{0}\right)=0\right) & =P\left(t_{0}<T_{1} \leq t_{1}<T_{1}+T_{2}, T_{1}<T_{1}^{\prime}\right) \\
& =\int_{t_{0}}^{t_{1}} P\left(T_{2}>t_{1}-x\right) P\left(T_{1}^{\prime}>x\right) \tau e^{-\tau x} d x \\
& =\int_{t_{0}}^{t_{1}} e^{-\tau\left(t_{1}-x\right)} e^{-\rho x} \tau e^{-\tau x} d x \\
& =\frac{\tau}{\rho} e^{-\tau t_{1}} \int_{t_{0}}^{t_{1}} \rho e^{-\rho x} d x=\frac{\tau}{\rho} e^{-\tau t_{1}}\left(e^{-\rho t_{0}}-e^{-\rho t_{1}}\right) .
\end{aligned}
$$

Analogously, we compute (by conditioning on $T_{1}, T_{2}$ for the second equality)

$$
\begin{aligned}
& P\left(f_{1}\left(t_{2}\right)=p+p q^{2}, f_{1}\left(t_{1}\right)=p, f_{1}\left(t_{0}\right)=0\right) \\
& \quad=P\left(t_{0}<T_{1} \leq t_{1}<T_{1}+T_{2} \leq t_{2}<T_{1}+T_{2}+T_{3}, T_{1}<T_{1}^{\prime} \leq T_{1}+T_{2}<T_{2}^{\prime}\right) \\
& \quad=\int_{t_{0}}^{t_{1}} \int_{t_{1}-x}^{t_{2}-x} P\left(T_{3}>t_{2}-x-y, T_{1}^{\prime} \in(x, x+y], T_{2}^{\prime}>x+y\right) \tau^{2} e^{-\tau x} e^{-\tau y} d y d x
\end{aligned}
$$

$$
\begin{aligned}
& =\tau^{2} \int_{t_{0}}^{t_{1}} \int_{t_{1}-x}^{t_{2}-x} e^{-\tau\left(t_{2}-x-y\right)} e^{-\rho x}\left(1-e^{-\rho y}\right) e^{-\rho(x+y)} e^{-\tau x} e^{-\tau y} d y d x \\
& \left.=\frac{\tau^{2}}{\rho} e^{-\tau t_{2}}\left[\rho^{-1}\left(e^{-\rho t_{1}}-e^{-\rho t_{2}}\right)\left(e^{-\rho t_{0}}-e^{-\rho t_{1}}\right)\right)-\frac{1}{2}\left(e^{-2 \rho t_{1}}-e^{-2 \rho t_{2}}\right)\left(t_{1}-t_{0}\right)\right]
\end{aligned}
$$

Taking the ratio shows that

$$
p\left(t_{2}, t_{1}, t_{0}\right)=\frac{\tau}{\rho} e^{-\tau\left(t_{2}-t_{1}\right)}\left(e^{-\rho t_{1}}-e^{-\rho t_{2}}\right)-\underbrace{\frac{\tau}{2} e^{-\tau\left(t_{2}-t_{1}\right)} \frac{e^{-2 \rho t_{1}}-e^{-2 \rho t_{2}}}{e^{-\rho t_{0}}-e^{-\rho t_{1}}}}_{\neq 0}\left(t_{1}-t_{0}\right)
$$

depends on $t_{0}$, so $f_{1}$ is not Markovian.
Remark 7. Our proof of Proposition 2 relies on the unique representation in $\mathcal{M}_{p}$. This means that it also holds true for all $p \in\left(0,2^{-1}\right)$ where each $x \in \mathcal{M}_{p}$ has a unique representation in $\mathcal{M}_{p}$, e.g. for all transcendental p. If the representation is not unique, Eq. (16) is still correct, but the right side of Eq. (17) does not show $P\left(f_{1}[1]=\right.$ $x)$. Instead, the latter shows the contribution to $P\left(f_{1}[1]=x\right)$ from the paths of $f_{1}$ which fulfil $C=j, B_{1}^{(j)}=b_{1}, \ldots, B_{j-1}^{(j)}=b_{j-1}$ (recall that $j, b_{1}, \ldots, b_{j-1}$ depend on the representation of $x$ ). Moreover, $P\left(f_{1}[1]=x\right)$ then is the sum over $P(C=j) P\left(B_{1}^{(j)}=b_{1}, \ldots, B_{j-1}^{(j)}=b_{j-1}\right)$ for the tuples $j, b_{1}, \ldots, b_{j-1}$ coming from the different representations of $x$ (the sets of paths are disjoint if the parameter sets $\left(j, b_{1}, \ldots, b_{j-1}\right)$ differ). Since the proof of our results on the Markov property of both $f_{1}$ and its jump chain also rely on the unique representation of $x$ (to read off which blocks merged when), the proof does not extend if $p$ does not allow a unique representation of $x$.

## 6 Example

We provide a concrete example showing that the random variables $\left(X_{k}\right)_{k \in \mathbb{N}}$ from Theorem 1 are, in general, neither independent nor identically distributed.

Choose $\Lambda=\delta_{\frac{1}{2}}$ and consider $f_{1}$ in the corresponding $\Lambda$-coalescent. Recall that $f_{1}[l]=x \in \mathcal{M}_{\frac{1}{2}}$ already fixes which singleton sets $\mathcal{S}_{k}$ are parts of the block of 1 at its $l$ th merger and which are not. First, assume $f_{1}[1]=X_{1}=\frac{5}{8}=\frac{1}{2}+\frac{1}{2^{3}} \in$ $\mathcal{M}_{\frac{1}{2}}$, which means that the coin of 1 comes up 'heads' for the first time at the third Poisson point and the block of 1 is $\mathcal{S}_{1} \cup \mathcal{S}_{3}$, while $\mathcal{S}_{2}$ is a block of its own (an event happening with probability $>0$ ). Assume further $f_{1}[2]=\frac{11}{16}=\frac{5}{8}+\frac{1}{16}$. This sets $X_{2}=\left(f_{1}[2]-f_{1}[1]\right) /\left(1-X_{1}\right)=\frac{1}{6} \notin \mathcal{M}_{\frac{1}{2}}$. We read off that the coin of the block of 1 also comes up 'heads' at the fourth collision, where the block of 1 merges with $\mathcal{S}_{4}$. We also see that the coin of the only other block $\mathcal{S}_{2}$ comes up 'tails'. We thus have, since we throw fair coins, $P\left(\left.X_{2}=\frac{1}{6} \right\rvert\, X_{1}=\frac{5}{8}\right)=P\left(\left.f_{1}[2]=\frac{11}{16} \right\rvert\, f_{1}[1]=\frac{5}{8}\right)=\frac{1}{4}$. Since $X_{1}=f_{1}[1] \in \mathcal{M}_{\frac{1}{2}}$ for any realisation, $X_{1}$ and $X_{2}$ have different distributions. To see also non-independence, consider $f_{1}[1]=X_{1}=\frac{1}{2}$ (coin of 1 comes up 'heads' at first coin toss, block of 1 is $\mathcal{S}_{1}$, occurs with probability $\left.\frac{1}{2}\right)$. In this case $P\left(X_{2}=\right.$ $\left.\frac{1}{6} \left\lvert\, X_{1}=\frac{1}{2}\right.\right)=0$, since $f_{1}[2]=X_{1}+\left(1-X_{1}\right) X_{2}=\frac{7}{12} \notin \mathcal{M}_{\frac{1}{2}}$.

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